# Package 'msm' 

September 27, 2021

## Version 1.6.9

Date 2021-09-26
Title Multi-State Markov and Hidden Markov Models in Continuous Time
Author Christopher Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)
Maintainer Christopher Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)
Description Functions for fitting continuous-time Markov and hidden Markov multi-state models to longitudinal data. Designed for processes observed at arbitrary times in continuous time (panel data) but some other observation schemes are supported. Both Markov transition rates and the hidden Markov output process can be modelled in terms of covariates, which may be constant or piecewise-constant in time.

License GPL (>=2)
Imports survival,mvtnorm,expm
Suggests mstate,minqa,doParallel,foreach,numDeriv,testthat,flexsurv
URL https://github.com/chjackson/msm
BugReports https://github.com/chjackson/msm/issues
LazyData yes
NeedsCompilation yes
Repository CRAN
Date/Publication 2021-09-27 09:50:02 UTC

## $R$ topics documented:

2phase ..... 3
aneur ..... 5
boot.msm ..... 6
bos ..... 9
cav ..... 10
cmodel.object ..... 11
coef.msm ..... 11
crudeinits.msm ..... 12
deltamethod ..... 13
draic.msm ..... 15
ecmodel.object ..... 18
efpt.msm ..... 18
ematrix.msm ..... 21
emodel.object ..... 22
fev ..... 23
hazard.msm ..... 24
hmm-dists ..... 25
hmmMV ..... 28
hmodel.object ..... 30
logLik.msm ..... 32
1rtest.msm ..... 33
MatrixExp ..... 33
medists ..... 35
model.frame.msm ..... 37
msm ..... 39
msm.form.qoutput ..... 52
msm.object ..... 53
msm.summary ..... 55
msm2Surv ..... 56
odds.msm ..... 58
paramdata.object ..... 59
pearson.msm ..... 61
pexp ..... 65
phasemeans.msm ..... 66
plot.msm ..... 67
plot.prevalence.msm ..... 68
plot.survfit.msm ..... 70
plotprog.msm ..... 72
pmatrix.msm ..... 73
pmatrix.piecewise.msm ..... 75
pnext.msm ..... 77
ppass.msm ..... 79
prevalence.msm ..... 81
print.msm ..... 84
printold.msm ..... 85
psor ..... 86
qcmodel.object ..... 87
qgeneric ..... 88
qmatrix.msm ..... 89
qmodel.object ..... 91
qratio.msm ..... 92
recreate.olddata ..... 93
scoreresid.msm ..... 94
sim.msm ..... 95
simfitted.msm ..... 96
simmulti.msm ..... 97
sojourn.msm ..... 99
statetable.msm ..... 101
surface.msm ..... 102
tnorm ..... 103
totlos.msm ..... 105
transient.msm ..... 108
updatepars.msm ..... 109
viterbi.msm ..... 109
Index ..... 111
2phase Coxian phase-type distribution with two phases

## Description

Density, distribution, quantile functions and other utilities for the Coxian phase-type distribution with two phases.

## Usage

```
d2phase(x, l1, mu1, mu2, log=FALSE)
p2phase(q, l1, mu1, mu2, lower.tail=TRUE, log.p=FALSE)
q2phase(p, l1, mu1, mu2, lower.tail=TRUE, log.p=FALSE)
r2phase(n, l1, mu1, mu2)
h2phase(x, l1, mu1, mu2, log=FALSE)
```


## Arguments

$x, q \quad$ vector of quantiles.
$p \quad$ vector of probabilities.
$n \quad$ number of observations. If length $(n)>1$, the length is taken to be the number required.
$11 \quad$ Intensity for transition between phase 1 and phase 2.
mu1 Intensity for transition from phase 1 to exit.
mu2 Intensity for transition from phase 2 to exit.
$\log \quad$ logical; if TRUE, return $\log$ density or $\log$ hazard.
log.p logical; if TRUE, probabilities p are given as $\log (\mathrm{p})$.
lower.tail logical; if TRUE (default), probabilities are $\mathrm{P}[\mathrm{X}<=\mathrm{x}]$, otherwise, $\mathrm{P}[\mathrm{X}>\mathrm{x}]$.

## Details

This is the distribution of the time to reach state 3 in a continuous-time Markov model with three states and transitions permitted from state 1 to state 2 (with intensity $\lambda_{1}$ ) state 1 to state 3 (intensity $\mu_{1}$ ) and state 2 to state 3 (intensity $\mu_{2}$ ). States 1 and 2 are the two "phases" and state 3 is the "exit" state.
The density is

$$
f\left(t \mid \lambda_{1}, \mu_{1}\right)=e^{-\left(\lambda_{1}+\mu_{1}\right) t}\left(\mu_{1}+\left(\lambda_{1}+\mu_{1}\right) \lambda_{1} t\right)
$$

if $\lambda_{1}+\mu_{1}=\mu_{2}$, and

$$
f\left(t \mid \lambda_{1}, \mu_{1}, \mu_{2}\right)=\frac{\left(\lambda_{1}+\mu_{1}\right) e^{-\left(\lambda_{1}+\mu_{1}\right) t}\left(\mu_{2}-\mu_{1}\right)+\mu_{2} \lambda_{1} e^{-\mu_{2} t}}{\lambda_{1}+\mu_{1}-\mu_{2}}
$$

otherwise. The distribution function is

$$
F\left(t \mid \lambda_{1}, \mu_{1}\right)=1-e^{-\left(\lambda_{1}+\mu_{1}\right) t}\left(1+\lambda_{1} t\right)
$$

if $\lambda_{1}+\mu_{1}=\mu_{2}$, and

$$
F\left(t \mid \lambda_{1}, \mu_{1}, \mu_{2}\right)=1-\frac{e^{-\left(\lambda_{1}+\mu_{1}\right) t}\left(\mu_{2}-\mu_{1}\right)+\lambda_{1} e^{-\mu_{2} t}}{\lambda_{1}+\mu_{1}-\mu_{2}}
$$

otherwise. Quantiles are calculated by numerically inverting the distribution function.
The mean is $\left(1+\lambda_{1} / \mu_{2}\right) /\left(\lambda_{1}+\mu_{1}\right)$.
The variance is $\left(2+2 \lambda_{1}\left(\lambda_{1}+\mu_{1}+\mu_{2}\right) / \mu_{2}^{2}-\left(1+\lambda_{1} / \mu_{2}\right)^{2}\right) /\left(\lambda_{1}+\mu_{1}\right)^{2}$.
If $\mu_{1}=\mu_{2}$ it reduces to an exponential distribution with rate $\mu_{1}$, and the parameter $\lambda_{1}$ is redundant. Or also if $\lambda_{1}=0$.
The hazard at $x=0$ is $\mu_{1}$, and smoothly increasing if $\mu_{1}<\mu_{2}$. If $\lambda_{1}+\mu_{1} \geq \mu_{2}$ it increases to an asymptote of $\mu_{2}$, and if $\lambda_{1}+\mu_{1} \leq \mu_{2}$ it increases to an asymptote of $\lambda_{1}+\mu_{1}$. The hazard is decreasing if $\mu_{1}>\mu_{2}$, to an asymptote of $\mu_{2}$.

## Value

d2phase gives the density, p2phase gives the distribution function, q2phase gives the quantile function, r2phase generates random deviates, and h2phase gives the hazard.

## Alternative parameterisation

An individual following this distribution can be seen as coming from a mixture of two populations:

1) "short stayers" whose mean sojourn time is $M_{1}=1 /\left(\lambda_{1}+\mu_{1}\right)$ and sojourn distribution is exponential with rate $\lambda_{1}+\mu_{1}$.
2) "long stayers" whose mean sojourn time $M_{2}=1 /\left(\lambda_{1}+\mu_{1}\right)+1 / \mu_{2}$ and sojourn distribution is the sum of two exponentials with rate $\lambda_{1}+\mu_{1}$ and $\mu_{2}$ respectively. The individual is a "long stayer" with probability $p=\lambda_{1} /\left(\lambda_{1}+\mu_{1}\right)$.

Thus a two-phase distribution can be more intuitively parameterised by the short and long stay means $M_{1}<M_{2}$ and the long stay probability $p$. Given these parameters, the transition intensities are $\lambda_{1}=p / M_{1}, \mu_{1}=(1-p) / M_{1}$, and $\mu_{2}=1 /\left(M_{2}-M_{1}\right)$. This can be useful for choosing intuitively reasonable initial values for procedures to fit these models to data.
The hazard is increasing at least if $M_{2}<2 M_{1}$, and also only if $\left(M_{2}-2 M_{1}\right) /\left(M_{2}-M_{1}\right)<p$.
For increasing hazards with $\lambda_{1}+\mu_{1} \leq \mu_{2}$, the maximum hazard ratio between any time $t$ and time 0 is $1 /(1-p)$.
For increasing hazards with $\lambda_{1}+\mu_{1} \geq \mu_{2}$, the maximum hazard ratio is $M_{1} /\left((1-p)\left(M_{2}-M_{1}\right)\right)$. This is the minimum hazard ratio for decreasing hazards.

## General phase-type distributions

This is a special case of the n-phase Coxian phase-type distribution, which in turn is a special case of the (general) phase-type distribution. The actuar R package implements a general n-phase distribution defined by the time to absorption of a general continuous-time Markov chain with a single absorbing state, where the process starts in one of the transient states with a given probability.

Author(s)
C. H. Jackson <chris.jackson@mrc-bsu. cam.ac.uk>

## References

C. Dutang, V. Goulet and M. Pigeon (2008). actuar: An R Package for Actuarial Science. Journal of Statistical Software, vol. 25, no. 7, 1-37. URL http://www.jstatsoft.org/v25/i07
aneur Aortic aneurysm progression data

## Description

This dataset contains longitudinal measurements of grades of aortic aneurysms, measured by ultrasound examination of the diameter of the aorta.

## Usage

aneur

## Format

A data frame containing 4337 rows, with each row corresponding to an ultrasound scan from one of 838 men over 65 years of age.

| ptnum | (numeric) | Patient identification number |
| ---: | :--- | :--- |
| age | (numeric) | Recipient age at examination (years) |
| diam | (numeric) | Aortic diameter |
| state | (numeric) | State of aneurysm. |

The states represent successive degrees of aneurysm severity, as indicated by the aortic diameter.

| State 1 | Aneurysm-free | $<30 \mathrm{~cm}$ |
| :--- | :--- | :--- |
| State 2 | Mild aneurysm | $30-44 \mathrm{~cm}$ |
| State 3 | Moderate aneurysm | $45-54 \mathrm{~cm}$ |
| State 4 | Severe aneurysm | $>55 \mathrm{~cm}$ |

683 of these men were aneurysm-free at age 65 and were re-screened every two years. The remaining men were aneurysmal at entry and had successive screens with frequency depending on the state of the aneurysm. Severe aneurysms are repaired by surgery.

## Source

The Chichester, U.K. randomised controlled trial of screening for abdominal aortic aneurysms by ultrasonography.

## References

Jackson, C.H., Sharples, L.D., Thompson, S.G. and Duffy, S.W. and Couto, E. Multi-state Markov models for disease progression with classification error. The Statistician, 52(2): 193-209 (2003)

Couto, E. and Duffy, S. W. and Ashton, H. A. and Walker, N. M. and Myles, J. P. and Scott, R. A. P. and Thompson, S. G. (2002) Probabilities of progression of aortic aneurysms: estimates and implications for screening policy Journal of Medical Screening 9(1):40-42

```
boot.msm Bootstrap resampling for multi-state models
```


## Description

Draw a number of bootstrap resamples, refit a msm model to the resamples, and calculate statistics on the refitted models.

## Usage

boot.msm(x, stat=pmatrix.msm, B=1000, file=NULL, cores=NULL)

## Arguments

x
stat

B

A fitted msm model, as output by msm.
A function to call on each refitted msm model. By default this is pmatrix.msm, returning the transition probability matrix in one time unit. If NULL then no function is computed.
Number of bootstrap resamples.

```
file Name of a file in which to save partial results after each replicate. This is
    saved using save and can be restored using load, producing an object called
    boot.list containing the partial results. Not supported when using parallel
    processing.
cores Number of processor cores to use for parallel processing. Requires the doPar-
    allel package to be installed. If not specified, parallel processing is not used. If
    cores is set to the string "default", the default methods of makeCluster (on
    Windows) or registerDoParallel (on Unix-like) are used.
```


## Details

The bootstrap datasets are computed by resampling independent transitions between pairs of states (for non-hidden models without censoring), or independent individual series (for hidden models or models with censoring). Therefore this approach doesn't work if, for example, the data for a HMM consist of a series of observations from just one individual, and is inaccurate for small numbers of independent transitions or individuals.
Confidence intervals or standard errors for the corresponding statistic can be calculated by summarising the returned list of $B$ replicated outputs. This is currently implemented for most the output functions qmatrix.msm, ematrix.msm, qratio.msm, pmatrix.msm, pmatrix.piecewise.msm, totlos.msm and prevalence.msm. For other outputs, users will have to write their own code to summarise the output of boot.msm.
Most of msm's output functions present confidence intervals based on asymptotic standard errors calculated from the Hessian. These are expected to be underestimates of the true standard errors (Cramer-Rao lower bound). Some of these functions use a further approximation, the delta method (see deltamethod) to obtain standard errors of transformed parameters. Bootstrapping should give a more accurate estimate of the uncertainty.
An alternative method which is less accurate though faster than bootstrapping, but more accurate than the delta method, is to draw a sample from the asymptotic multivariate normal distribution implied by the maximum likelihood estimates (and covariance matrix), and summarise the transformed estimates. See pmatrix.msm.
All objects used in the original call to msm which produced $x$, such as the qmatrix, should be in the working environment, or else boot.msm will produce an "object not found" error. This enables boot.msm to refit the original model to the replicate datasets. However there is currently a limitation. In the original call to msm, the "formula" argument should be specified directly, as, for example,
msm(state $\sim$ time, data $=\ldots$ )
and not, for example,
form $=$ data\$state $\sim$ data\$time
msm(formula=form, data = ...)
otherwise boot.msm will be unable to draw the replicate datasets.
boot.msm will also fail with an incomprehensible error if the original call to msm used a useddefined object whose name is the same as a built-in R object, or an object in any other loaded package. For example, if you have called a $Q$ matrix $q$, when $q()$ is the built-in function for quitting R .

If stat is NULL, then B different msm model objects will be stored in memory. This is unadvisable, as msm objects tend to be large, since they contain the original data used for the msm fit, so this will be wasteful of memory.

To specify more than one statistic, write a function consisting of a list of different function calls, for example,

```
stat = function(x) list (pmatrix.msm(x,t=1), pmatrix.msm(x,t=2))
```


## Value

A list with $B$ components, containing the result of calling function stat on each of the refitted models. If stat is NULL, then each component just contains the refitted model. If one of the B model fits was unsuccessful and resulted in an error, then the corresponding list component will contain the error message.

## Author(s)

C.H.Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## References

Efron, B. and Tibshirani, R.J. (1993) An Introduction to the Bootstrap, Chapman and Hall.

## See Also

qmatrix.msm, qratio.msm, sojourn.msm, ematrix.msm, pmatrix.msm, pmatrix.piecewise.msm, totlos.msm, prevalence.msm.

## Examples

```
## Not run:
    ## Psoriatic arthritis example
    data(psor)
    psor.q <- rbind(c(0,0.1,0,0),c(0,0,0.1,0),c(0,0,0,0.1),c(0,0,0,0))
    psor.msm <- msm(state ~ months, subject=ptnum, data=psor, qmatrix =
        psor.q, covariates = ~ollwsdrt+hieffusn,
        constraint = list(hieffusn=c(1,1,1),ollwsdrt=c(1,1,2)),
        control = list(REPORT=1,trace=2), method="BFGS")
    ## Bootstrap the baseline transition intensity matrix. This will take a long time.
    q.list <- boot.msm(psor.msm, function(x)x$Qmatrices$baseline)
    ## Manipulate the resulting list of matrices to calculate bootstrap standard errors.
    apply(array(unlist(q.list), dim=c(4,4,5)), c(1,2), sd)
    ## Similarly calculate a bootstrap 95% confidence interval
    apply(array(unlist(q.list), dim=c(4,4,5)), c(1, 2),
            function(x)quantile(x, c(0.025, 0.975)))
    ## Bootstrap standard errors are larger than the asymptotic standard
    ## errors calculated from the Hessian
    psor.msm$QmatricesSE$baseline
## End(Not run)
```


## Description

A dataset containing histories of bronchiolitis obliterans syndrome (BOS) from lung transplant recipients. BOS is a chronic decline in lung function, often observed after lung transplantation. The condition is classified into four stages of severity: none, mild, moderate and severe.

## Usage

bos

## Format

A data frame containing 638 rows, grouped by patient, including histories of 204 patients. The first observation for each patient is defined to be stage 1 , no BOS, at six months after transplant. Subsequent observations denote the entry times into stages $2,3,4$, representing mild, moderate and severe BOS respectively, and stage 5, representing death.

| ptnum | (numeric) | Patient identification number |
| ---: | :--- | :--- |
| time | (numeric) | Months after transplant |
| state | (numeric) | BOS state entered at this time |

## Details

The entry time of each patient into each stage of BOS was estimated by clinicians, based on their history of lung function measurements and acute rejection and infection episodes. BOS is only assumed to occur beyond six months after transplant. In the first six months the function of each patient's new lung stabilises. Subsequently BOS is diagnosed by comparing the lung function against the "baseline" value.
The objects bos 3 and bos 4 contain the same data, but with mild/moderate/severe combined, and moderate/severe combined, to give 3 and 4-state representations respectively.

## Source

Papworth Hospital, U.K.

## References

Heng. D. et al. (1998). Bronchiolitis Obliterans Syndrome: Incidence, Natural History, Prognosis, and Risk Factors. Journal of Heart and Lung Transplantation 17(12)1255-1263.

## Description

A series of approximately yearly angiographic examinations of heart transplant recipients. The state at each time is a grade of cardiac allograft vasculopathy (CAV), a deterioration of the arterial walls.

## Usage

cav

## Format

A data frame containing 2846 rows. There are 622 patients, the rows are grouped by patient number and ordered by years after transplant, with each row representing an examination and containing additional covariates.

```
        PTNUM (numeric) Patient identification number
            age (numeric) Recipient age at examination (years)
        years (numeric) Examination time (years after transplant)
            dage (numeric) Age of heart donor (years)
            sex (numeric) sex (0=male, 1=female)
            pdiag (factor) Primary diagnosis (reason for transplant)
                            IHD=ischaemic heart disease, IDC=idiopathic dilated cardiomyopathy.
    cumrej (numeric) Cumulative number of acute rejection episodes
    state (numeric) State at the examination.
                                    State 1 represents no CAV, state 2 is mild/moderate CAV
                                    and state 3 is severe CAV. State 4 indicates death.
firstobs (numeric) 0= record represents an angiogram or date of death.
                            1 = record represents transplant (patient's first observation)
statemax (numeric) Maximum observed state so far for this patient (added in version 1.5.1)
```


## Source

Papworth Hospital, U.K.

## References

Sharples, L.D. and Jackson, C.H. and Parameshwar, J. and Wallwork, J. and Large, S.R. (2003). Diagnostic accuracy of coronary angiopathy and risk factors for post-heart-transplant cardiac allograft vasculopathy. Transplantation 76(4):679-82

## Description

A list giving information about censored states, their labels in the data and what true states they represent.

## Value

| ncens | The number of distinct values used for censored observations in the state data <br> supplied to msm. |
| :--- | :--- |
| censor | A vector of length ncens, giving the labels used for censored states in the data. |
| states | A vector obtained by unlist()ing a list with ncens elements, each giving the <br> set of true states that an observation with this label could be. |
| index | Index into states for the first state corresponding to each censor, plus an extra <br> length(states) +1. |

## See Also

msm.object.
coef.msm Extract model coefficients

## Description

Extract the estimated log transition intensities and the corresponding linear effects of each covariate.

## Usage

\#\# S3 method for class 'msm'
coef(object, ...)

## Arguments

object
. . .

A fitted multi-state model object, as returned by msm.
(unused) further arguments passed to or from other methods.

## Value

If there is no misclassification, coef.msm returns a list of matrices. The first component, labelled logbaseline, is a matrix containing the estimated transition intensities on the $\log$ scale with any covariates fixed at their means in the data. Each remaining component is a matrix giving the linear effects of the labelled covariate on the matrix of $\log$ intensities.

For misclassification models, coef.msm returns a list of lists. The first component, Qmatrices, is a list of matrices as described in the previous paragraph. The additional component Ematrices is a list of similar format containing the logit-misclassification probabilities and any estimated covariate effects.

## Author(s)

C. H. Jackson <chris.jackson@mrc-bsu. cam.ac.uk>

## See Also

msm
crudeinits.msm Calculate crude initial values for transition intensities

## Description

Calculates crude initial values for transition intensities by assuming that the data represent the exact transition times of the Markov process.

## Usage

crudeinits.msm(formula, subject, qmatrix, data=NULL, censor=NULL, censor.states=NULL)

## Arguments

formula A formula giving the vectors containing the observed states and the corresponding observation times. For example,
state ~ time
Observed states should be in the set $1, \ldots, n$, where $n$ is the number of states. Note hidden Markov models are not supported by this function.
subject Vector of subject identification numbers for the data specified by formula. If missing, then all observations are assumed to be on the same subject. These must be sorted so that all observations on the same subject are adjacent.
qmatrix Matrix of indicators for the allowed transitions. An initial value will be estimated for each value of qmatrix that is greater than zero. Transitions are taken as disallowed for each entry of qmatrix that is 0 .
data An optional data frame in which the variables represented by subject and state can be found.


#### Abstract

censor A state, or vector of states, which indicates censoring. See msm. censor.states Specifies the underlying states which censored observations can represent. See msm.


## Details

Suppose we want a crude estimate of the transition intensity $q_{r s}$ from state $r$ to state $s$. If we observe $n_{r s}$ transitions from state $r$ to state $s$, and a total of $n_{r}$ transitions from state $r$, then $q_{r s} / q_{r r}$ can be estimated by $n_{r s} / n_{r}$. Then, given a total of $T_{r}$ years spent in state $r$, the mean sojourn time $1 / q_{r r}$ can be estimated as $T_{r} / n_{r}$. Thus, $n_{r s} / T_{r}$ is a crude estimate of $q_{r s}$.
If the data do represent the exact transition times of the Markov process, then these are the exact maximum likelihood estimates.

Observed transitions which are incompatible with the given qmatrix are ignored. Censored states are ignored.

## Value

The estimated transition intensity matrix. This can be used as the qmatrix argument to msm.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

statetable.msm

## Examples

```
data(cav)
twoway4.q <- rbind(c(-0.5, 0.25, 0, 0.25), c(0.166, -0.498, 0.166, 0.166),
c(0, 0.25, -0.5, 0.25), c(0, 0, 0, 0))
statetable.msm(state, PTNUM, data=cav)
crudeinits.msm(state ~ years, PTNUM, data=cav, qmatrix=twoway4.q)
```

deltamethod The delta method

## Description

Delta method for approximating the standard error of a transformation $g(X)$ of a random variable $X=\left(x_{1}, x_{2}, \ldots\right)$, given estimates of the mean and covariance matrix of $X$.

## Usage

deltamethod(g, mean, cov, ses=TRUE)

## Arguments

g
A formula representing the transformation. The variables must be labelled $\times 1, \times 2, \ldots$ For example,
$\sim 1 /(x 1+x 2)$
If the transformation returns a vector, then a list of formulae representing $\left(g_{1}, g_{2}, \ldots\right)$ can be provided, for example
list( $\sim x 1+x 2, \sim x 1 /(x 1+x 2))$
mean The estimated mean of $X$
cov The estimated covariance matrix of $X$
ses If TRUE, then the standard errors of $g_{1}(X), g_{2}(X), \ldots$ are returned. Otherwise the covariance matrix of $g(X)$ is returned.

## Details

The delta method expands a differentiable function of a random variable about its mean, usually with a first-order Taylor approximation, and then takes the variance. For example, an approximation to the covariance matrix of $g(X)$ is given by

$$
\operatorname{Cov}(g(X))=g^{\prime}(\mu) \operatorname{Cov}(X)\left[g^{\prime}(\mu)\right]^{T}
$$

where $\mu$ is an estimate of the mean of $X$. This function uses symbolic differentiation via deriv.
A limitation of this function is that variables created by the user are not visible within the formula g. To work around this, it is necessary to build the formula as a string, using functions such as sprintf, then to convert the string to a formula using as. formula. See the example below.

If you can spare the computational time, bootstrapping is a more accurate method of calculating confidence intervals or standard errors for transformations of parameters. See boot.msm. Simulation from the asymptotic distribution of the MLEs (see e.g. Mandel 2013) is also a convenient alternative.

## Value

A vector containing the standard errors of $g_{1}(X), g_{2}(X), \ldots$ or a matrix containing the covariance of $g(X)$.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## References

Oehlert, G. W. (1992) A note on the delta method. American Statistician 46(1).
Mandel, M. (2013) Simulation based confidence intervals for functions with complicated derivatives. The American Statistician 67(2):76-81.

## Examples

```
## Simple linear regression, E(y) = alpha + beta x
x <- 1:100
y <- rnorm(100, 4*x, 5)
toy.lm <- lm(y ~ x)
estmean <- coef(toy.lm)
estvar <- summary(toy.lm)$cov.unscaled * summary(toy.lm)$sigma^2
## Estimate of (1 / (alphahat + betahat))
1 / (estmean[1] + estmean[2])
## Approximate standard error
deltamethod (~ 1 / (x1 + x2), estmean, estvar)
## We have a variable z we would like to use within the formula.
z <- 1
## deltamethod (~ z / (x1 + x2), estmean, estvar) will not work.
## Instead, build up the formula as a string, and convert to a formula.
form <- sprintf("~ %f / (x1 + x2)", z)
form
deltamethod(as.formula(form), estmean, estvar)
```

draic.msm Criteria for comparing two multi-state models with nested state spaces

## Description

A modification of Akaike's information criterion, and a leave-one-out likelihood cross-validation criterion, for comparing the predictive ability of two Markov multi-state models with nested state spaces. This is evaluated based on the restricted or aggregated data which the models have in common.

Note that standard AIC can be computed for one or more fitted msm models $x, y, \ldots$ using AIC ( $x, y, \ldots$ ), and this can be used to compare models fitted to the same data. draic.msm and drlcv.msm are designed for models fitted to data with differently-aggregated state spaces.

## Usage

```
draic.msm(msm.full, msm.coarse, likelihood.only=FALSE,
    information=c("expected","observed"), tl=0.95)
drlcv.msm(msm.full, msm.coarse, tl=0.95, cores=NULL,
            verbose=TRUE,outfile=NULL)
```


## Arguments

msm. full Model on the bigger state space.

| msm.coarse | Model on the smaller state space. |
| :---: | :---: |
|  | The two models must both be non-hidden Markov models without censored states. |
|  | The two models must be fitted to the same datasets, except that the state space of the coarse model must be an aggregated version of the state space of the full model. That is, every state in the full dataset must correspond to a unique state in the coarse dataset. For example, for the full state variable $c(1,1,2,2,3,4)$, the corresponding coarse states could be $c(1,1,2,2,2,3)$, but not $c(1,2,3,4,4,4)$ |
|  | The structure of allowed transitions in the coarse model must also be a collapsed version of the big model structure, but no check is currently made for this in the code. |
|  | To use these functions, all objects which were used in the calls to fit msm. full and msm. coarse must be in the working environment, for example, datasets and definitions of transition matrices. |
| likelihood.only |  |
|  | Don't calculate Hessians and trace term (DRAIC). |
| information | Use observed or expected information in the DRAIC trace term. Expected is the default, and much faster, though is only available for models fitted to pure panel data (all obstype=1 in the call to msm, thus not exact transition times or exact death times) |
| tl | Width of symmetric tracking interval, by default 0.95 for a $95 \%$ interval. |
| cores | Number of processor cores to use in drlcv for cross-validation by parallel processing. Requires the doParallel package to be installed. If not specified, parallel processing is not used. If cores is set to the string "default", the default methods of makeCluster (on Windows) or registerDoParallel (on Unixlike) are used. |
| verbose | Print intermediate results of each iteration of cross-validation to the console while running. May not work with parallel processing. |
| outfile | Output file to print intermediate results of cross-validation. Useful to track execution speed when using parallel processing, where output to the console may not work. |

## Details

The difference in restricted AIC (Liquet and Commenges, 2011), as computed by this function, is defined as

$$
D_{R A I C}=l\left(\gamma_{n} \mid \mathbf{x}^{\prime \prime}\right)-l\left(\theta_{n} \mid \mathbf{x}^{\prime \prime}\right)+\operatorname{trace}\left(J\left(\theta_{n} \mid \mathbf{x}^{\prime \prime}\right) J\left(\theta_{n} \mid \mathbf{x}\right)^{-1}-J\left(\gamma_{n} \mid \mathbf{x}^{\prime \prime}\right) J\left(\gamma_{n} \mid \mathbf{x}^{\prime}\right)^{-1}\right)
$$

where $\gamma$ and $\theta$ are the maximum likelihood estimates of the smaller and bigger models, fitted to the smaller and bigger data, respectively.
$l\left(\gamma_{n} \mid x^{\prime \prime}\right)$ represents the likelihood of the simpler model evaluated on the restricted data.
$l\left(\theta_{n} \mid x^{\prime \prime}\right)$ represents the likelihood of the complex model evaluated on the restricted data. This is a hidden Markov model, with a misclassification matrix and initial state occupancy probabilities as described by Thom et al (2014).
$J()$ are the corresponding (expected or observed, as specified by the user) information matrices.
$\mathbf{x}$ is the expanded data, to which the bigger model was originally fitted, and $\mathbf{x}^{\prime}$ is the data to which the smaller model was originally fitted. $\mathrm{x}^{\prime \prime}$ is the restricted data which the two models have in common. $\mathrm{x}^{\prime \prime}=\mathrm{x}^{\prime}$ in this implementation, so the models are nested.

The difference in likelihood cross-validatory criteria (Liquet and Commenges, 2011) is defined as

$$
D_{R L C V}=1 / n \sum_{i=1}^{n} \log \left(h_{X^{\prime \prime}}\left(x_{i}^{\prime \prime} \mid \gamma_{-i}\right) / g_{X^{\prime \prime}}\left(x_{i}^{\prime \prime} \mid \theta_{-i}\right)\right)
$$

where $\gamma_{-i}$ and $\theta_{-i}$ are the maximum likelihood estimates from the smaller and bigger models fitted to datasets with subject $i$ left out, $g()$ and $h()$ are the densities of the corresponding models, and $x_{i}^{\prime \prime}$ is the restricted data from subject $i$.

Tracking intervals are analogous to confidence intervals, but not strictly the same, since the quantity which D_RAIC aims to estimate, the difference in expected Kullback-Leibler discrepancy for predicting a replicate dataset, depends on the sample size. See the references.

Positive values for these criteria indicate the coarse model is preferred, while negative values indicate the full model is preferred.

## Value

A list containing $D_{R A I C}$ (draic.msm) or $D_{R L C V}(\mathrm{drlcv} . \mathrm{msm})$, its component terms, and tracking intervals.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk), H. H. Z. Thom <howard. thom@bristol.ac.uk>

## References

Thom, H. and Jackson, C. and Commenges, D. and Sharples, L. (2015) State selection in multistate models with application to quality of life in psoriatic arthritis. Statistics In Medicine 34(16) 2381 2480.

Liquet, B. and Commenges D. (2011) Choice of estimators based on different observations: Modified AIC and LCV criteria. Scandinavian Journal of Statistics; 38:268-287.

## See Also

logLik.msm

```
ecmodel.object
```

Developer documentation: model for covariates on misclassification probabilities

## Description

A list representing the model for covariates on misclassification probabilities.

## Value

npars Number of covariate effect parameters. This is defined as the number of covariates on misclassification (with factors expanded as contrasts) multiplied by the number of allowed misclassifications in the model.
ndpars $\quad$ Number of distinct covariate effect parameters, as npars, but after any equality constraints have been applied.
ncovs Number of covariates on misclassification, with factors expanded as contrasts.
constr List of equality constraints on these covariate effects, as supplied in the miscconstraint argument to msm.
covlabels Names/labels of these covariates in the model matrix (see model.matrix.msm).
inits Initial values for these covariate effects, as a vector formed from the misccovinits list supplied to msm.
covmeans Means of these covariates in the data (excluding data not required to fit the model, such as observations with missing data in other elements or subjects' last observations). This includes means of $0 / 1$ factor contrasts as well as continuous covariates (for historic reasons, which may not be sensible).

## See Also

msm.object.
efpt.msm Expected first passage time

## Description

Expected time until first reaching a particular state or set of states in a Markov model.

## Usage

```
efpt.msm(x=NULL, qmatrix=NULL, tostate, start="all", covariates="mean",
    ci=c("none","normal","bootstrap"), cl=0.95, B=1000,
    cores=NULL, ...)
```


## Arguments

$\left.\begin{array}{ll}\mathrm{x} & \text { A fitted multi-state model, as returned by msm. } \\ \text { qmatrix } & \text { Instead of x, you can simply supply a transition intensity matrix in qmatrix. } \\ \text { tostate } & \begin{array}{l}\text { State, or set of states supplied as a vector, for which to estimate the first passage } \\ \text { time into. Can be integer, or character matched to the row names of the Q matrix. }\end{array} \\ \text { start } & \begin{array}{l}\text { Starting state (integer). By default (start="all"), this will return a vector of } \\ \text { expected passage times from each state in turn. }\end{array} \\ & \text { Alternatively, this can be used to obtain the expected first passage time from a } \\ & \text { set of states, rather than single states. To achieve this, state is set to a vector of } \\ \text { weights, with length equal to the number of states in the model. These weights } \\ \text { should be proportional to the probability of starting in each of the states in the } \\ \text { desired set, so that weights of zero are supplied for other states. The function }\end{array}\right\}$

## Details

The expected first passage times from each of a set of states $\mathbf{i}$ to to the remaining set of states $\overline{\mathbf{i}}$ in the state space, for a model with transition intensity matrix $Q$, are

$$
-Q_{i, i}^{-1} 1
$$

where $\mathbf{1}$ is a vector of ones, and $Q_{\mathbf{i}, \mathbf{i}}$ is the square subset of $Q$ pertaining to states $\mathbf{i}$.
It is equal to the sum of mean sojourn times for all states between the "from" and "to" states in a unidirectional model. If there is non-zero chance of reaching an absorbing state before reaching tostate, then it is infinite. It is trivially zero if the "from" state equals tostate.

This function currently only handles time-homogeneous Markov models. For time-inhomogeneous models it will assume that $Q$ equals the average intensity matrix over all times and observed covariates. Simulation might be used to handle time dependence.

Note this is the expectation of first passage time, and the confidence intervals are CIs for this mean, not predictive intervals for the first passage time. The full distribution of the first passage time to a set of states can be obtained by setting the rows of the intensity matrix $Q$ corresponding to that set of states to zero to make a model where those states are absorbing. The corresponding transition probability matrix $\operatorname{Exp}(Q t)$ then gives the probabilities of having hit or passed that state by a time $t$ (see the example below). This is implemented in ppass.msm.

## Value

A vector of expected first passage times, or "hitting times", from each state to the desired state.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## References

Norris, J. R. (1997) Markov Chains. Cambridge University Press.

## See Also

sojourn.msm, totlos.msm, boot.msm.

## Examples

```
twoway4.q <- rbind(c(-0.5, 0.25, 0, 0.25), c(0.166, -0.498, 0.166, 0.166),
    c(0, 0.25, -0.5, 0.25), c(0, 0, 0, 0))
efpt.msm(qmatrix=twoway4.q, tostate=3)
# given in state 1, expected time to reaching state 3 is infinite
# since may die (state 4) before entering state 3
# If we remove the death state from the model, EFPTs become finite
Q <- twoway4.q[1:3,1:3]; diag(Q) <- 0; diag(Q) <- -rowSums(Q)
efpt.msm(qmatrix=Q, tostate=3)
# Suppose we cannot die or regress while in state 2, can only go to state 3
Q <- twoway4.q; Q[2,4] <- Q[2,1] <- 0; diag(Q) <- 0; diag(Q) <- -rowSums(Q)
efpt.msm(qmatrix=Q, tostate=3)
# The expected time from 2 to 3 now equals the mean sojourn time in 2.
-1/Q[2, 2]
# Calculate cumulative distribution of the first passage time
# into state 3 for the following three-state model
Q <- twoway4.q[1:3,1:3]; diag(Q) <- 0; diag(Q) <- -rowSums(Q)
# Firstly form a model where the desired hitting state is absorbing
Q[3,] <- 0
MatrixExp(Q, t=10)[,3]
ppass.msm(qmatrix=Q, tot=10)
# Given in state 1 at time 0, P(hit 3 by time 10) = 0.479
MatrixExp(Q, t=50)[,3] # P(hit 3 by time 50) = 0.98
ppass.msm(qmatrix=Q, tot=50)
```


## ematrix.msm Misclassification probability matrix

## Description

Extract the estimated misclassification probability matrix, and corresponding confidence intervals, from a fitted multi-state model at a given set of covariate values.

## Usage

ematrix.msm(x, covariates="mean", ci=c("delta","normal","bootstrap","none"), $\mathrm{cl}=0.95$, $\mathrm{B}=1000$, cores=NULL)

## Arguments

$x \quad$ A fitted multi-state model, as returned by msm.
covariates The covariate values for which to estimate the misclassification probability matrix. This can either be:
the string "mean", denoting the means of the covariates in the data (this is the default),
the number 0 , indicating that all the covariates should be set to zero,
or a list of values, with optional names. For example
list $(60,1)$
where the order of the list follows the order of the covariates originally given in the model formula, or a named list,
list (age $=60$, sex $=1$ )
ci If "delta" (the default) then confidence intervals are calculated by the delta method, or by simple transformation of the Hessian in the very simplest cases.
If "normal", then calculate a confidence interval by simulating B random vectors from the asymptotic multivariate normal distribution implied by the maximum likelihood estimates (and covariance matrix) of the multinomial-logittransformed misclassification probabilities and covariate effects, then transforming back.
If "bootstrap" then calculate a confidence interval by non-parametric bootstrap refitting. This is 1-2 orders of magnitude slower than the "normal" method, but is expected to be more accurate. See boot.msm for more details of bootstrapping in $\mathbf{~ m s m}$.
cl Width of the symmetric confidence interval to present. Defaults to 0.95.
B Number of bootstrap replicates, or number of normal simulations from the distribution of the MLEs
cores Number of cores to use for bootstrapping using parallel processing. See boot.msm for more details.

## Details

Misclassification probabilities and covariate effects are estimated on the multinomial-logit scale by msm. A covariance matrix is estimated from the Hessian of the maximised log-likelihood. From these, the delta method can be used to obtain standard errors of the probabilities on the natural scale at arbitrary covariate values. Confidence intervals are estimated by assuming normality on the multinomial-logit scale.

## Value

A list with components:
estimate Estimated misclassification probability matrix. The rows correspond to true states, and columns observed states.

SE Corresponding approximate standard errors.
L Lower confidence limits.
U Upper confidence limits.

Or if ci="none", then ematrix.msm just returns the estimated misclassification probability matrix.
The default print method for objects returned by ematrix.msm presents estimates and confidence limits. To present estimates and standard errors, do something like

```
ematrix.msm(x)[c("estimates","SE")]
```


## Author(s)

C. H. Jackson <chris.jackson@mrc-bsu. cam.ac.uk>

## See Also

qmatrix.msm

```
emodel.object
```


## Description

A list giving information about the misclassifications assumed in a multi-state model fitted with the ematrix argument of msm. Returned in a fitted msm model object. This information is converted internally to a hmodel object (see hmodel. object) for use in likelihood computations.

| Value |  |
| :--- | :--- |
| nstates |  |
| npars |  |
| imatrix | Number of states (same as qmodel\$nstates). <br> Number of allowed misclassifications, equal to sum(imatrix). <br> Indicator matrix for allowed misclassifications. This has $(r, s)$ entry 1 if mis- <br> classification of true state $r$ as observed state $s$ is possible. diagonal entries are <br> arbitrarily set to 0. |
| ematrix | Matrix of initial values for the misclassification probabilities, supplied as the <br> ematrix argument of msm. |
| inits | Vector of these initial values, reading across rows of qmatrix and excluding the <br> diagonal and disallowed transitions. <br> Indicators for equality constraints on baseline misclassification probabilities, <br> taken from the econstraint argument to msm, and mapped if necessary to the |
| ndpars | set $(1,2,3, \ldots)$. <br> Number of distinct misclassification probabilities, after applying equality con- <br> straints. |
| nipars | Number of initial state occupancy probabilities being estimated. This is zero if <br> est.initprobs=FALSE, otherwise equal to the number of states. |
| initprobs | Initial state occupancy probabilities, as supplied to msm (initial values before <br> estimation, if est. initprobs=TRUE.) |
| est.initprobsAre initial state occupancy probabilities estimated (TRUE or FALSE), as supplied <br> in the est.initprobs argument of msm. |  |

## See Also

msm. object,qmodel.object, hmodel. object.
fev FEV1 measurements from lung transplant recipients

## Description

A series of measurements of the forced expiratory volume in one second (FEV1) from lung transplant recipients, from six months onwards after their transplant.

## Usage

fev

## Format

A data frame containing 5896 rows. There are 204 patients, the rows are grouped by patient number and ordered by days after transplant. Each row represents an examination and containing an additional covariate.

| ptnum | (numeric) | Patient identification number. |
| ---: | :--- | :--- |
| days | (numeric) | Examination time (days after transplant). |
| fev | (numeric) | Percentage of baseline FEV1. A code of 999 indicates the patient's date of death. |
| acute | (numeric) | $0 / 1$ indicator for whether the patient suffered an acute infection or rejection |
|  |  | within 14 days of the visit. |

## Details

A baseline "normal" FEV1 for each individual is calculated using measurements from the first six months after transplant. After six months, as presented in this dataset, FEV1 is expressed as a percentage of the baseline value.
FEV1 is monitored to diagnose bronchiolitis obliterans syndrome (BOS), a long-term lung function decline, thought to be a form of chronic rejection. Acute rejections and infections also affect the lung function in the short term.

## Source

Papworth Hospital, U.K.

## References

Jackson, C.H. and Sharples, L.D. Hidden Markov models for the onset and progression of bronchiolitis obliterans syndrome in lung transplant recipients Statistics in Medicine, 21(1): 113-128 (2002).
hazard.msm Calculate tables of hazard ratios for covariates on transition intensities

## Description

Hazard ratios are computed by exponentiating the estimated covariate effects on the log-transition intensities. This function is called by summary.msm.

## Usage

hazard.msm(x, hazard.scale $=1$, cl $=0.95)$

## Arguments

x
Output from msm representing a fitted multi-state model.

Vector with same elements as number of covariates on transition rates. Corresponds to the increase in each covariate used to calculate its hazard ratio. Defaults to all 1 .
cl Width of the symmetric confidence interval to present. Defaults to 0.95.

## Value

A list of tables containing hazard ratio estimates, one table for each covariate. Each table has three columns, containing the hazard ratio, and an approximate upper and lower confidence limit respectively (assuming normality on the log scale), for each Markov chain transition intensity.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

msm, summary.msm, odds.msm
hmm-dists Hidden Markov model constructors

## Description

These functions are used to specify the distribution of the response conditionally on the underlying state in a hidden Markov model. A list of these function calls, with one component for each state, should be used for the hmodel argument to msm. The initial values for the parameters of the distribution should be given as arguments. Note the initial values should be supplied as literal values supplying them as variables is currently not supported.

## Usage

```
hmmCat(prob, basecat)
hmmIdent(x)
hmmUnif(lower, upper)
hmmNorm(mean, sd)
hmmLNorm(meanlog, sdlog)
hmmExp(rate)
hmmGamma(shape, rate)
hmmWeibull(shape, scale)
hmmPois(rate)
hmmBinom(size, prob)
hmmTNorm(mean, sd, lower, upper)
hmmMETNorm(mean, sd, lower, upper, sderr, meanerr=0)
hmmMEUnif(lower, upper, sderr, meanerr=0)
hmmNBinom(disp, prob)
hmmBetaBinom(size, meanp, sdp)
hmmBeta(shape1, shape2)
hmmT(mean, scale,df)
```

| Arguments |  |
| :---: | :--- |
| prob | (hmmCat) Vector of probabilities of observing category $1,2, \ldots$, length (prob) <br> respectively. Or the probability governing a binomial or negative binomial dis- <br> tribution. |
| basecat | (hmmCat) Category which is considered to be the "baseline", so that during esti- <br> mation, the probabilities are parameterised as probabilities relative to this base- <br> line category. By default, the category with the greatest probability is used as <br> the baseline. |
| (hmmIdent) Code in the data which denotes the exactly-observed state. |  |
| (hmmNorm, hmmLNorm, hmmTNorm) Mean defining a Normal, or truncated Normal |  |

## Details

hmmCat represents a categorical response distribution on the set $1,2, \ldots$, length (prob). The Markov model with misclassification is an example of this type of model. The categories in this case are (some subset of) the underlying states.
The hmmIdent distribution is used for underlying states which are observed exactly without error. For hidden Markov models with multiple outcomes, (see hmmMV), the outcome in the data which takes the special hmmIdent value must be the first of the multiple outcomes.
hmmUnif, hmmNorm, hmmLNorm, hmmExp, hmmGamma, hmmWeibull, hmmPois, hmmBinom, hmmTNorm, hmmNBinom and hmmBeta represent Uniform, Normal, log-Normal, exponential, Gamma, Weibull, Poisson, Binomial, truncated Normal, negative binomial and beta distributions, respectively, with parameterisations the same as the default parameterisations in the corresponding base R distribution functions.
hmmT is the Student t distribution with general mean $\mu$, scale $\sigma$ and degrees of freedom df . The variance is $\sigma^{2} d f /(d f+2)$. Note the t distribution in base $\mathrm{R} d t$ is a standardised one with mean 0 and scale 1 . These allow any positive (integer or non-integer) df. By default, all three parameters, including df, are estimated when fitting a hidden Markov model, but in practice, df might need to be fixed for identifiability - this can be done using the fixedpars argument to msm.
The hmmMETNorm and hmmMEUnif distributions are truncated Normal and Uniform distributions, but with additional Normal measurement error on the response. These are generalisations of the distributions proposed by Satten and Longini (1996) for modelling the progression of CD4 cell counts in monitoring HIV disease. See medists for density, distribution, quantile and random generation functions for these distributions. See also tnorm for density, distribution, quantile and random generation functions for the truncated Normal distribution.
See the PDF manual 'msm-manual.pdf' in the 'doc' subdirectory for algebraic definitions of all these distributions. New hidden Markov model response distributions can be added to msm by following the instructions in Section 2.17.1.
Parameters which can be modelled in terms of covariates, on the scale of a link function, are as follows.

| PARAMETER NAME | LINK FUNCTION |
| :--- | :--- |
| mean | identity |
| meanlog | identity |
| rate | $\log$ |
| scale | $\log$ |
| meanerr | identity |
| meanp | logit |
| prob | logit or multinomial logit |

Parameters basecat, lower, upper, size, meanerr are fixed at their initial values. All other parameters are estimated while fitting the hidden Markov model, unless the appropriate fixedpars argument is supplied to msm .

For categorical response distributions (hmmCat) the outcome probabilities initialized to zero are fixed at zero, and the probability corresponding to basecat is fixed to one minus the sum of the remaining probabilities. These remaining probabilities are estimated, and can be modelled in terms of covariates via multinomial logistic regression (relative to basecat).

## Value

Each function returns an object of class hmodel, which is a list containing information about the model. The only component which may be useful to end users is $r$, a function of one argument $n$ which returns a random sample of size n from the given distribution.

## Author(s)

C. H. Jackson <chris.jackson@mrc-bsu. cam.ac.uk>

## References

Satten, G.A. and Longini, I.M. Markov chains with measurement error: estimating the 'true' course of a marker of the progression of human immunodeficiency virus disease (with discussion) Applied Statistics 45(3): 275-309 (1996).

Jackson, C.H. and Sharples, L.D. Hidden Markov models for the onset and progresison of bronchiolitis obliterans syndrome in lung transplant recipients Statistics in Medicine, 21(1): 113-128 (2002).

Jackson, C.H., Sharples, L.D., Thompson, S.G. and Duffy, S.W. and Couto, E. Multi-state Markov models for disease progression with classification error. The Statistician, 52(2): 193-209 (2003).

## See Also

msm

## hmmMV

Multivariate hidden Markov models

## Description

Constructor for a a multivariate hidden Markov model (HMM) where each of the $n$ variables observed at the same time has a (potentially different) standard univariate distribution conditionally on the underlying state. The n outcomes are independent conditionally on the hidden state.
If a particular state in a HMM has such an outcome distribution, then a call to hmmMV is supplied as the corresponding element of the hmodel argument to msm. See Example 2 below.
A multivariate HMM where multiple outcomes at the same time are generated from the same distribution is specified in the same way as the corresponding univariate model, so that hmmMV is not required. The outcome data are simply supplied as a matrix instead of a vector. See Example 1 below.
The outcome data for such models are supplied as a matrix, with number of columns equal to the maximum number of arguments supplied to the hmmMV calls for each state. If some but not all of the variables are missing (NA) at a particular time, then the observed data at that time still contribute to the likelihood. The missing data are assumed to be missing at random. The Viterbi algorithm may be used to predict the missing values given the fitted model and the observed data.
Typically the outcome model for each state will be from the same family or set of families, but with different parameters. Theoretically, different numbers of distributions may be supplied for different
states. If a particular state has fewer outcomes than the maximum, then the data for that state are taken from the first columns of the response data matrix. However this is not likely to be a useful model, since the number of observations will probably give information about the underlying state, violating the missing at random assumption.
Models with outcomes that are dependent conditionally on the hidden state (e.g. correlated multivariate normal observations) are not currently supported.

## Usage

hmmMV (...)

## Arguments

... The number of arguments supplied should equal the maximum number of observations made at one time. Each argument represents the univariate distribution of that outcome conditionally on the hidden state, and should be the result of calling a univariate hidden Markov model constructor (see hmm-dists).

## Value

A list of objects, each of class hmmdist as returned by the univariate HMM constructors documented in hmm-dists. The whole list has class hmmMVdist, which inherits from hmmdist.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## References

Jackson, C. H., Su, L., Gladman, D. D. and Farewell, V. T. (2015) On modelling minimal disease activity. Arthritis Care and Research (early view).

## See Also

```
hmm-dists,msm
```


## Examples

```
## Simulate data from a Markov model
nsubj <- 30; nobspt <- 5
sim.df <- data.frame(subject = rep(1:nsubj, each=nobspt),
                                    time = seq(0, 20, length=nobspt))
set.seed(1)
two.q <- rbind(c(-0.1, 0.1), c(0, 0))
dat <- simmulti.msm(sim.df[,1:2], qmatrix=two.q, drop.absorb=FALSE)
### EXAMPLE 1
## Generate two observations at each time from the same outcome
## distribution:
## Bin(40, 0.1) for state 1, Bin(40, 0.5) for state 2
dat$obs1[dat$state==1] <- rbinom(sum(dat$state==1), 40, 0.1)
```

```
dat$obs2[dat$state==1] <- rbinom(sum(dat$state==1), 40, 0.1)
dat$obs1[dat$state==2] <- rbinom(sum(dat$state==2), 40, 0.5)
dat$obs2[dat$state==2] <- rbinom(sum(dat$state==2), 40, 0.5)
dat$obs <- cbind(obs1 = dat$obs1, obs2 = dat$obs2)
## Fitted model should approximately recover true parameters
msm(obs ~ time, subject=subject, data=dat, qmatrix=two.q,
    hmodel = list(hmmBinom(size=40, prob=0.2),
            hmmBinom(size=40, prob=0.2)))
### EXAMPLE 2
## Generate two observations at each time from different
## outcome distributions:
## Bin(40, 0.1) and Bin(40, 0.2) for state 1,
dat$obs1 <- dat$obs2 <- NA
dat$obs1[dat$state==1] <- rbinom(sum(dat$state==1), 40, 0.1)
dat$obs2[dat$state==1] <- rbinom(sum(dat$state==1), 40, 0.2)
## Bin(40, 0.5) and Bin(40, 0.6) for state 2
dat$obs1[dat$state==2] <- rbinom(sum(dat$state==2), 40, 0.6)
dat$obs2[dat$state==2] <- rbinom(sum(dat$state==2), 40, 0.5)
dat$obs <- cbind(obs1 = dat$obs1, obs2 = dat$obs2)
## Fitted model should approximately recover true parameters
msm(obs ~ time, subject=subject, data=dat, qmatrix=two.q,
    hmodel = list(hmmMV(hmmBinom(size=40, prob=0.3),
                    hmmBinom(size=40, prob=0.3)),
        hmmMV(hmmBinom(size=40, prob=0.3),
                        hmmBinom(size=40, prob=0.3))),
        control=list(maxit=10000))
```

hmodel. object Developer documentation: hidden Markov model structure object

## Description

A list giving information about the models for the outcome data conditionally on the states of a hidden Markov model. Used in internal computations, and returned in a fitted msm model object.

## Value

| hidden | TRUE for hidden Markov models, FALSE otherwise. |
| :--- | :--- |
| nstates | Number of states, the same as qmodel\$nstates. |
| fitted | TRUE if the parameter values in pars are the maximum likelihood estimates, <br> FALSE if they are the initial values. |
| models | The outcome distribution for each hidden state. A vector of length nstates <br> whose $r$ th entry is the index of the state $r$ outcome distributions in the vector of <br> supported distributions. The vector of supported distributions is given in full by <br> msm: $:=$. msm. HMODELS: the first few are 1 for categorical outcome, 2 for identity, <br>  <br>  for uniform and 4 for normal. |


| labels | String identifying each distribution in models. |
| :---: | :---: |
| npars | Vector of length nstates giving the number of parameters in each outcome distribution, excluding covariate effects. |
| nipars | Number of initial state occupancy probabilities being estimated. This is zero if est.initprobs=FALSE, otherwise equal to the number of states. |
| totpar | Total number |
| pars | A vector of length totpars, made from concatenating a list of length nstates whose $r$ th component is vector of the parameters for the state $r$ outcome distribution. |
| plabs | List with the names of the parameters in pars. |
| parstate | A vector of length totpars, whose $i$ th element is the state corresponding to the $i$ th parameter. |
| firstpar | A vector of length nstates giving the index in pars of the first parameter for each state. |
| locpars | Index in pars of parameters which can have covariates on them. |
| initprobs | Initial state occupancy probabilities, as supplied to msm (initial values before estimation, if est.initprobs=TRUE.) |
| est.initprobs | Are initial state occupancy probabilities estimated (TRUE or FALSE), as supplied in the est.initprobs argument of msm. |
| ncovs | Number of covariate effects per parameter in pars, with, e.g. factor contrasts expanded. |
| coveffect | Vector of covariate effects, of length sum(ncovs). |
| covlabels | Labels of these effects. |
| coveffstate | Vector indicating state corresponding to each element of coveffect. |
| ncoveffs | Number of covariate effects on HMM outcomes, equal to sum(ncovs). |
| nicovs | Vector of length nstates-1 giving the number of covariate effects on each initial state occupancy probability (log relative to the baseline probability). |
| icoveffect | Vector of length sum(nicovs) giving covariate effects on initial state occupancy probabilities. |
| nicoveffs | Number of covariate effects on initial state occupancy probabilities, equal to sum(nicovs). |
| constr | Constraints on (baseline) hidden Markov model outcome parameters, as supplied in the hconstraint argument of msm, excluding covariate effects, converted to a vector and mapped to the set $1,2,3, \ldots$ if necessary. |
| covconstr | Vector of constraints on covariate effects in hidden Markov outcome models, as supplied in the hconstraint argument of msm, excluding baseline parameters, converted to a vector and mapped to the set $1,2,3, \ldots$ if necessary. |
| ranges | Matrix of range restrictions for HMM parameters, including those given to the hranges argument to msm. |
| foundse | TRUE if standard errors are available for the estimates. |
| initpmat | Matrix of initial state occupancy probabilities with one row for each subject (estimated if est.initprobs=TRUE). |
| ci | Confidence intervals for baseline HMM outcome parameters. |
| covci | Confidence intervals for covariate effects in HMM outcome models. |

## See Also

msm.object,qmodel.object, emodel.object.
logLik.msm Extract model log-likelihood

## Description

Extract the log-likelihood and the number of parameters of a model fitted with msm.

## Usage

\#\# S3 method for class 'msm'
$\operatorname{logLik}$ (object, by.subject=FALSE, ...)

## Arguments

object A fitted multi-state model object, as returned by msm.
by.subject
Return vector of subject-specific log-likelihoods, which should sum to the total log-likelihood.
.. (unused) further arguments passed to or from other methods.

## Value

The log-likelihood of the model represented by 'object' evaluated at the maximum likelihood estimates.

Akaike's information criterion can also be computed using AIC (object).

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

msm,lrtest.msm.

```
    lrtest.msm Likelihood ratio test
```


## Description

Likelihood ratio test between two or more fitted multi-state models

## Usage

lrtest.msm(...)

## Arguments

$$
\begin{array}{ll}
\ldots & \text { Two or more fitted multi-state models, as returned by msm, ordered by increasing } \\
\text { numbers of parameters. }
\end{array}
$$

## Value

A matrix with three columns, giving the likelihood ratio statistic, difference in degrees of freedom and the chi-squared p-value for a comparison of the first model supplied with each subsequent model.

## Warning

The comparison between models will only be valid if they are fitted to the same dataset. This may be a problem if there are missing values and R's default of 'na.action = na.omit' is used.
The likelihood ratio statistic only has the indicated chi-squared distribution if the models are nested. An alternative for comparing non-nested models is Akaike's information criterion. This can be computed for one or more fitted msm models $x, y, \ldots$ using AIC ( $x, y, \ldots$ ).

See Also
logLik.msm,msm
MatrixExp Matrix exponential

## Description

Calculates the exponential of a square matrix.

## Usage

MatrixExp(mat, $\mathrm{t}=1$, method=NULL, ...)

## Arguments

mat A square matrix
t
method

An optional scaling factor for mat.
Under the default of NULL, this simply wraps the expm function from the expm package. This is recommended. Options to expm can be supplied to MatrixExp, including method.
Otherwise, for backwards compatibility, the following options, which use code in the msm package, are available: "pade" for a Pade approximation method, "series" for the power series approximation, or "analytic" for the analytic formulae for simpler Markov model intensity matrices (see below). These options are only used if mat has repeated eigenvalues, thus the usual eigen-decomposition method cannot be used.
... Arguments to pass to expm.

## Details

See the expm documentation for details of the algorithms it uses.
Generally the exponential $E$ of a square matrix $M$ can often be calculated as

$$
E=U \exp (D) U^{-1}
$$

where $D$ is a diagonal matrix with the eigenvalues of $M$ on the diagonal, $\exp (D)$ is a diagonal matrix with the exponentiated eigenvalues of $M$ on the diagonal, and $U$ is a matrix whose columns are the eigenvectors of $M$.

This method of calculation is used if "pade" or "series" is supplied but $M$ has distinct eigenvalues. I If $M$ has repeated eigenvalues, then its eigenvector matrix may be non-invertible. In this case, the matrix exponential is calculated using the Pade approximation defined by Moler and van Loan (2003), or the less robust power series approximation,

$$
\exp (M)=I+M+M^{2} / 2+M^{3} / 3!+M^{4} / 4!+\ldots
$$

For a continuous-time homogeneous Markov process with transition intensity matrix $Q$, the probability of occupying state $s$ at time $u+t$ conditional on occupying state $r$ at time $u$ is given by the $(r, s)$ entry of the matrix $\exp (t Q)$.
If mat is a valid transition intensity matrix for a continuous-time Markov model (i.e. diagonal entries non-positive, off-diagonal entries non-negative, rows sum to zero), then for certain simpler model structures, there are analytic formulae for the individual entries of the exponential of mat. These structures are listed in the PDF manual and the formulae are coded in the msm source file src/analyticp.c. These formulae are only used if method="analytic". This is more efficient, but it is not the default in MatrixExp because the code is not robust to extreme values. However it is the default when calculating likelihoods for models fitted by msm.
The implementation of the Pade approximation used by method="pade" was taken from JAGS by Martyn Plummer (https://mcmc-jags.sourceforge.io).
medists

## Value

The exponentiated matrix $\exp (m a t)$. Or, if $t$ is a vector of length 2 or more, an array of exponentiated matrices.

## References

Cox, D. R. and Miller, H. D. The theory of stochastic processes, Chapman and Hall, London (1965)
Moler, C and van Loan, C (2003). Nineteen dubious ways to compute the exponential of a matrix, twenty-five years later. SIAM Review 45, 3-49.

```
medists Measurement error distributions
```


## Description

Truncated Normal and Uniform distributions, where the response is also subject to a Normally distributed measurement error.

## Usage

```
dmenorm(x, mean=0, sd=1, lower=-Inf, upper=Inf, sderr=0, meanerr=0,
    log = FALSE)
pmenorm(q, mean=0, sd=1, lower=-Inf, upper=Inf, sderr=0, meanerr=0,
    lower.tail = TRUE, log.p = FALSE)
qmenorm(p, mean=0, sd=1, lower=-Inf, upper=Inf, sderr=0, meanerr=0,
    lower.tail = TRUE, log.p = FALSE)
rmenorm(n, mean=0, sd=1, lower=-Inf, upper=Inf, sderr=0, meanerr=0)
dmeunif(x, lower=0, upper=1, sderr=0, meanerr=0, log = FALSE)
pmeunif(q, lower=0, upper=1, sderr=0, meanerr=0, lower.tail = TRUE,
    log.p = FALSE)
qmeunif(p, lower=0, upper=1, sderr=0, meanerr=0, lower.tail = TRUE,
        log.p = FALSE)
rmeunif(n, lower=0, upper=1, sderr=0, meanerr=0)
```


## Arguments

$x, q \quad$ vector of quantiles.
$p \quad$ vector of probabilities.
$\mathrm{n} \quad$ number of observations. If length $(\mathrm{n})>1$, the length is taken to be the number required.
mean vector of means.
sd vector of standard deviations.
lower lower truncation point.
upper upper truncation point.
sderr Standard deviation of measurement error distribution.
meanerr Optional shift for the measurement error distribution.
$\log , \log . \mathrm{p} \quad \operatorname{logical}$; if TRUE, probabilities $p$ are given as $\log (p)$, or $\log$ density is returned.
lower. tail logical; if TRUE (default), probabilities are $P[X<=x]$, otherwise, $P[X>x]$.

## Details

The normal distribution with measurement error has density

$$
\frac{\Phi\left(u, \mu_{2}, \sigma_{3}\right)-\Phi\left(l, \mu_{2}, \sigma_{3}\right)}{\Phi\left(u, \mu_{0}, \sigma_{0}\right)-\Phi\left(l, \mu_{0}, \sigma_{0}\right)} \phi\left(x, \mu_{0}+\mu_{\epsilon}, \sigma_{2}\right)
$$

where

$$
\begin{gathered}
\sigma_{2}^{2}=\sigma_{0}^{2}+\sigma_{\epsilon}^{2} \\
\sigma_{3}=\sigma_{0} \sigma_{\epsilon} / \sigma_{2} \\
\mu_{2}=\left(x-\mu_{\epsilon}\right) \sigma_{0}^{2}+\mu_{0} \sigma_{\epsilon}^{2}
\end{gathered}
$$

$\mu_{0}$ is the mean of the original Normal distribution before truncation, $\sigma_{0}$ is the corresponding standard deviation, $u$ is the upper truncation point, $l$ is the lower truncation point, $\sigma_{\epsilon}$ is the standard deviation of the additional measurement error, $\mu_{\epsilon}$ is the mean of the measurement error (usually 0 ).
$\phi(x)$ is the density of the corresponding normal distribution, and $\Phi(x)$ is the distribution function of the corresponding normal distribution.
The uniform distribution with measurement error has density

$$
\left(\Phi\left(x, \mu_{\epsilon}+l, \sigma_{\epsilon}\right)-\Phi\left(x, \mu_{\epsilon}+u, \sigma_{\epsilon}\right)\right) /(u-l)
$$

These are calculated from the original truncated Normal or Uniform density functions $f(. \mid \mu, \sigma, l, u)$ as

$$
\int f(y \mid \mu, \sigma, l, u) \phi\left(x, y+\mu_{\epsilon}, \sigma_{\epsilon}\right) d y
$$

If sderr and meanerr are not specified they assume the default values of 0 , representing no measurement error variance, and no constant shift in the measurement error, respectively.

Therefore, for example with no other arguments, dmenorm( $x$ ), is simply equivalent to dtnorm ( $x$ ), which in turn is equivalent to dnorm ( $x$ ).
These distributions were used by Satten and Longini (1996) for CD4 cell counts conditionally on hidden Markov states of HIV infection, and later by Jackson and Sharples (2002) for FEV1 measurements conditionally on states of chronic lung transplant rejection.

These distribution functions are just provided for convenience, and are not optimised for numerical accuracy or speed. To fit a hidden Markov model with these response distributions, use a hmmMETNorm or hmmMEUnif constructor. See the hmm-dists help page for further details.

## Value

dmenorm, dmeunif give the density, pmenorm, pmeunif give the distribution function, qmenorm, qmeunif give the quantile function, and rmenorm, rmeunif generate random deviates, for the Normal and Uniform versions respectively.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## References

Satten, G.A. and Longini, I.M. Markov chains with measurement error: estimating the 'true' course of a marker of the progression of human immunodeficiency virus disease (with discussion) Applied Statistics 45(3): 275-309 (1996)
Jackson, C.H. and Sharples, L.D. Hidden Markov models for the onset and progression of bronchiolitis obliterans syndrome in lung transplant recipients Statistics in Medicine, 21(1): 113-128 (2002).

## See Also

dnorm, dunif, dtnorm

## Examples

```
## what does the distribution look like?
x <- seq(50, 90, by=1)
plot(x, dnorm(x, 70, 10), type="l", ylim=c(0,0.06)) ## standard Normal
lines(x, dtnorm(x, 70, 10, 60, 80), type="l") ## truncated Normal
## truncated Normal with small measurement error
lines(x, dmenorm(x, 70, 10, 60, 80, sderr=3), type="l")
```

    model. frame.msm Extract original data from msm objects.
    
## Description

Extract the data from a multi-state model fitted with msm.

## Usage

```
## S3 method for class 'msm'
model.frame(formula, agg=FALSE, ...)
## S3 method for class 'msm'
model.matrix(object, model="intens", state=1, ...)
```


## Arguments

| formula | A fitted multi-state model object, as returned by msm. |
| :--- | :--- |
| agg | Return the model frame in the efficient aggregated form used to calculate the <br> likelihood internally for non-hidden Markov models. This has one row for each <br> unique combination of from-state, to-state, time lag, covariate value and obser- <br> vation type. The variable named " (nocc)" counts how many observations of <br> that combination there are in the original data. |
| object | A fitted multi-state model object, as returned by msm. |
| model | "intens" to return the design matrix for covariates on intensities, "misc" for <br> misclassification probabilities, "hmm" for a general hidden Markov model, and <br> "inits" for initial state probabilities in hidden Markov models. |
| state | State corresponding to the required covariate design matrix in a hidden Markov <br> model. |
| $\ldots$ | Further arguments (not used). |

## Value

model.frame returns a data frame with all the original variables used for the model fit, with any missing data removed (see na.action in msm). The state, time, subject, obstype and obstrue variables are named "(state)", "(time)", "(subject)", "(obstype)" and "(obstrue)" respectively (note the brackets). A variable called " (obs)" is the observation number from the original data before any missing data were dropped. The variable "(pcomb)" is used for computing the likelihood for hidden Markov models, and identifies which distinct time difference, obstype and covariate values (thus which distinct interval transition probability matrix) each observation corresponds to.
The model frame object has some other useful attributes, including "usernames" giving the user's original names for these variables (used for model refitting, e.g. in bootstrapping or cross validation) and "covnames" identifying which ones are covariates.
model.matrix returns a design matrix for a part of the model that includes covariates. The required part is indicated by the "model" argument.

For time-inhomogeneous models fitted with "pci", these datasets will have imputed observations at each time change point, indicated where the variable " (pci.imp)" in the model frame is 1 . The model matrix for intensities will have factor contrasts for the timeperiod covariate.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

msm, model.frame, model.matrix.

## Description

Fit a continuous-time Markov or hidden Markov multi-state model by maximum likelihood. Observations of the process can be made at arbitrary times, or the exact times of transition between states can be known. Covariates can be fitted to the Markov chain transition intensities or to the hidden Markov observation process.

## Usage

```
    msm ( formula, subject=NULL, data = list(), qmatrix, gen.inits = FALSE,
        ematrix=NULL, hmodel=NULL, obstype=NULL, obstrue=NULL,
        covariates = NULL, covinits = NULL, constraint = NULL,
        misccovariates = NULL, misccovinits = NULL, miscconstraint = NULL,
        hcovariates = NULL, hcovinits = NULL, hconstraint = NULL, hranges=NULL,
        qconstraint=NULL, econstraint=NULL, initprobs = NULL,
        est.initprobs=FALSE, initcovariates = NULL, initcovinits = NULL,
        deathexact = NULL, death=NULL, exacttimes = FALSE, censor=NULL,
        censor.states=NULL, pci=NULL, phase.states=NULL, phase.inits=NULL,
        cl = 0.95, fixedpars = NULL, center=TRUE,
        opt.method="optim", hessian=NULL, use.deriv=TRUE,
        use.expm=TRUE, analyticp=TRUE, na.action=na.omit, ... )
```


## Arguments

formula A formula giving the vectors containing the observed states and the corresponding observation times. For example,
state ~ time
Observed states should be numeric variables in the set $1, \ldots, n$, where $n$ is the number of states. Factors are allowed only if their levels are called " 1 ", . . , "n". The times can indicate different types of observation scheme, so be careful to choose the correct obstype.
For hidden Markov models, state refers to the outcome variable, which need not be a discrete state. It may also be a matrix, giving multiple observations at each time (see hmmMV).
subject Vector of subject identification numbers for the data specified by formula. If missing, then all observations are assumed to be on the same subject. These must be sorted so that all observations on the same subject are adjacent.
data Optional data frame in which to interpret the variables supplied in formula, subject, covariates, misccovariates, hcovariates, obstype and obstrue.
qmatrix Matrix which indicates the allowed transitions in the continuous-time Markov chain, and optionally also the initial values of those transitions. If an instantaneous transition is not allowed from state $r$ to state $s$, then qmatrix should have $(r, s)$ entry 0 , otherwise it should be non-zero.

If supplying initial values yourself, then the non-zero entries should be those values. If using gen.inits=TRUE then the non-zero entries can be anything you like (conventionally 1). Any diagonal entry of qmatrix is ignored, as it is constrained to be equal to minus the sum of the rest of the row.
For example,

```
rbind(c(0,0.1,0.01),c(0.1,0,0.2),c(0,0,0))
```

represents a 'health - disease - death' model, with initial transition intensities 0.1 from health to disease, 0.01 from health to death, 0.1 from disease to health, and 0.2 from disease to death.
If the states represent ordered levels of severity of a disease, then this matrix should usually only allow transitions between adjacent states. For example, if someone was observed in state 1 ("mild") at their first observation, followed by state 3 ("severe") at their second observation, they are assumed to have passed through state 2 ("moderate") in between, and the 1,3 entry of qmatrix should be zero.
The initial intensities given here are with any covariates set to their means in the data (or set to zero, if center $=$ FALSE). If any intensities are constrained to be equal using qconstraint, then the initial value is taken from the first of these (reading across rows).
gen.inits If TRUE, then initial values for the transition intensities are generated automatically using the method in crudeinits.msm. The non-zero entries of the supplied qmatrix are assumed to indicate the allowed transitions of the model. This is not available for hidden Markov models, including models with misclassified states.
ematrix If misclassification between states is to be modelled, this should be a matrix of initial values for the misclassification probabilities. The rows represent underlying states, and the columns represent observed states. If an observation of state $s$ is not possible when the subject occupies underlying state $r$, then ematrix should have $(r, s)$ entry 0 . Otherwise ematrix should have $(r, s)$ entry corresponding to the probability of observing $s$ conditionally on occupying true state $r$. The diagonal of ematrix is ignored, as rows are constrained to sum to 1 . For example,

```
rbind(c(0,0.1,0),c(0.1,0,0.1),c(0,0.1,0))
```

represents a model in which misclassifications are only permitted between adjacent states.
If any probabilities are constrained to be equal using econstraint, then the initial value is taken from the first of these (reading across rows).
For an alternative way of specifying misclassification models, see hmodel.
hmodel Specification of the hidden Markov model (HMM). This should be a list of return values from HMM constructor functions. Each element of the list corresponds to the outcome model conditionally on the corresponding underlying state. Univariate constructors are described in thehmm-dists help page. These
may also be grouped together to specify a multivariate HMM with a set of conditionally independent univariate outcomes at each time, as described in hmmMV.
For example, consider a three-state hidden Markov model. Suppose the observations in underlying state 1 are generated from a Normal distribution with mean 100 and standard deviation 16, while observations in underlying state 2 are Normal with mean 54 and standard deviation 18. Observations in state 3, representing death, are exactly observed, and coded as 999 in the data. This model is specified as
hmodel = list(hmmNorm(mean=100, sd=16), hmmNorm(mean=54, sd=18),hmmIdent (999))
The mean and standard deviation parameters are estimated starting from these initial values. If multiple parameters are constrained to be equal using hconstraint, then the initial value is taken from the value given on the first occasion that parameter appears in hmodel.
See the hmm-dists help page for details of the constructor functions for each univariate distribution.
A misclassification model, that is, a hidden Markov model where the outcomes are misclassified observations of the underlying states, can either be specified using a list of hmmCat or hmmIdent objects, or by using an ematrix.
For example,

```
ematrix = rbind(c(0,0.1,0,0),c(0.1,0,0.1,0),c(0,0.1,0,0),c(0,0,0,0)
)
```

is equivalent to
hmodel $=\operatorname{list}(\operatorname{hmmCat}(\operatorname{prob}=c(0.9,0.1,0,0)), \operatorname{hmmCat}(\operatorname{prob}=c(0.1,0.8,0.1,0))$, hmmCat $(p r o b=$
obstype A vector specifying the observation scheme for each row of the data. This can be included in the data frame data along with the state, time, subject IDs and covariates. Its elements should be either 1, 2 or 3, meaning as follows:

1 An observation of the process at an arbitrary time (a "snapshot" of the process, or "panel-observed" data). The states are unknown between observation times.
2 An exact transition time, with the state at the previous observation retained until the current observation. An observation may represent a transition to a different state or a repeated observation of the same state (e.g. at the end of follow-up). Note that if all transition times are known, more flexible models could be fitted with packages other than msm - see the note under exacttimes.
Note also that if the previous state was censored using censor, for example known only to be state 1 or state 2 , then obstype 2 means that either state 1 is retained or state 2 is retained until the current observation - this does not allow for a change of state in the middle of the observation interval.
3 An exact transition time, but the state at the instant before entering this state is unknown. A common example is death times in studies of chronic diseases.

If obstype is not specified, this defaults to all 1. If obstype is a single number, all observations are assumed to be of this type. The obstype value for the first observation from each subject is not used.
This is a generalisation of the deathexact and exacttimes arguments to allow different schemes per observation. obstype overrides both deathexact and exacttimes.
exacttimes=TRUE specifies that all observations are of obstype 2 .
deathexact $=$ death. states specifies that all observations of death. states are of type 3. deathexact $=$ TRUE specifies that all observations in the final absorbing state are of type 3 .
obstrue In misclassification models specified with ematrix, obstrue is a vector of logicals (TRUE or FALSE) or numerics ( 1 or 0 ) specifying which observations (TRUE, $1)$ are observations of the underlying state without error, and which (FALSE, 0) are realisations of a hidden Markov model.
In HMMs specified with hmodel, where the hidden state is known at some times, if obstrue is supplied it is assumed to contain the actual true state data. Elements of obstrue at times when the hidden state is unknown are set to NA. This allows the information from HMM outcomes generated conditionally on the known state to be included in the model, thus improving the estimation of the HMM outcome distributions.
In HMMs where there are also censored states, obstrue should be set to 1 for observed states which are censored but not misclassified.
covariates A formula or a list of formulae representing the covariates on the transition intensities via a log-linear model. If a single formula is supplied, like covariates $=\sim$ age + sex + treatment
then these covariates are assumed to apply to all intensities. If a named list is supplied, then this defines a potentially different model for each named intensity. For example,
covariates = list("1-2" = ~ age, "2-3" = ~ age + treatment)
specifies an age effect on the state 1 - state 2 transition, additive age and treatment effects on the state 2 - state 3 transition, but no covariates on any other transitions that are allowed by the qmatrix.
If covariates are time dependent, they are assumed to be constant in between the times they are observed, and the transition probability between a pair of times $(t 1, t 2)$ is assumed to depend on the covariate value at $t 1$.
covinits Initial values for log-linear effects of covariates on the transition intensities. This should be a named list with each element corresponding to a covariate. A single element contains the initial values for that covariate on each transition intensity, reading across the rows in order. For a pair of effects constrained to be equal, the initial value for the first of the two effects is used.
For example, for a model with the above qmatrix and age and sex covariates, the following initialises all covariate effects to zero apart from the age effect on the 2-1 transition, and the sex effect on the 1-3 transition. covinits = list (sex=c $(0,0,0.1,0)$, age $=c(0,0.1,0,0))$
For factor covariates, name each level by concatenating the name of the covariate with the level name, quoting if necessary. For example, for a covariate agegroup with three levels $0-15,15-60,60-$, use something like

miscconstraint A list of one vector for each named covariate on misclassification probabilities. The vector indicates which covariate effects on misclassification probabilities are constrained to be equal, analogously to constraint. Only used if the model is specified using ematrix.
hcovariates List of formulae the same length as hmodel, defining any covariates governing the hidden Markov outcome models. The covariates operate on a suitably linktransformed linear scale, for example, log scale for a Poisson outcome model. If there are no covariates for a certain hidden state, then insert a NULL in the corresponding place in the list. For example, hcovariates = list(~acute + age, ~acute, NULL).
hcovinits Initial values for the hidden Markov model covariate effects. A list of the same length as hcovariates. Each element is a vector with initial values for the effect of each covariate on that state. For example, the above hcovariates can be initialised with hcovariates $=\operatorname{list}(c(-8,0),-8, N U L L)$. Initial values must be given for all or no covariates, if none are given these are all set to zero. The initial value given in the hmodel constructor function for the corresponding baseline parameter is interpreted as the value of that parameter with any covariates fixed to their means in the data. If multiple effects are constrained to be equal using hconstraint, then the initial value is taken from the first of the multiple initial values supplied.
hconstraint A named list. Each element is a vector of constraints on the named hidden Markov model parameter. The vector has length equal to the number of times that class of parameter appears in the whole model.
For example consider the three-state hidden Markov model described above, with normally-distributed outcomes for states 1 and 2 . To constrain the outcome variance to be equal for states 1 and 2 , and to also constrain the effect of acute on the outcome mean to be equal for states 1 and 2 , specify
hconstraint $=$ list ( $s d=c(1,1)$, acute $=c(1,1)$ )
Note this excludes initial state occupancy probabilities and covariate effects on those probabilities, which cannot be constrained.
hranges Range constraints for hidden Markov model parameters. Supplied as a named list, with each element corresponding to the named hidden Markov model parameter. This element is itself a list with two elements, vectors named "lower" and "upper". These vectors each have length equal to the number of times that class of parameter appears in the whole model, and give the corresponding mininum amd maximum allowable values for that parameter. Maximum likelihood estimation is performed with these parameters constrained in these ranges (through a log or logit-type transformation). Lower bounds of -Inf and upper bounds of Inf can be given if the parameter is unbounded above or below. For example, in the three-state model above, to constrain the mean for state 1 to be between 0 and 6 , and the mean of state 2 to be between 7 and 12, supply
hranges=list(mean=list(lower=c ( 0,7 ), upper=c $(6,12)$ ))
These default to the natural ranges, e.g. the positive real line for variance parameters, and $[0,1]$ for probabilities. Therefore hranges need not be specified for such parameters unless an even stricter constraint is desired. If only one limit is supplied for a parameter, only the first occurrence of that parameter is constrained.

Initial values should be strictly within any ranges, and not on the range boundary, otherwise optimisation will fail with a "non-finite value" error.
qconstraint A vector of indicators specifying which baseline transition intensities are equal. For example,
qconstraint $=c(1,2,3,3)$
constrains the third and fourth intensities to be equal, in a model with four allowed instantaneous transitions. When there are covariates on the intensities and center=TRUE (the default), qconstraint is applied to the intensities with covariates taking the values of the means in the data. When center=FALSE, qconstraint is applied to the intensities with covariates set to zero.
econstraint A similar vector of indicators specifying which baseline misclassification probabilities are constrained to be equal. Only used if the model is specified using ematrix, rather than hmodel.
initprobs Only used in hidden Markov models. Underlying state occupancy probabilities at each subject's first observation. Can either be a vector of nstates elements with common probabilities to all subjects, or a nsubjects by nstates matrix of subject-specific probabilities. This refers to observations after missing data and subjects with only one observation have been excluded.
If these are estimated (see est. initprobs), then this represents an initial value, and defaults to equal probability for each state. Otherwise this defaults to $c(1, r e p(0$, nstates -1$)$ ), that is, in state 1 with a probability of 1 . Scaled to sum to 1 if necessary. The state 1 occupancy probability should be non-zero.
est.initprobs Only used in hidden Markov models. If TRUE, then the underlying state occupancy probabilities at the first observation will be estimated, starting from a vector of initial values supplied in the initprobs argument. Structural zeroes are allowed: if any of these initial values are zero they will be fixed at zero during optimisation, even if est.initprobs=TRUE, and no covariate effects on them are estimated. The exception is state 1 , which should have non-zero occupancy probability.
Note that the free parameters during this estimation exclude the state 1 occupancy probability, which is fixed at one minus the sum of the other probabilities.
initcovariates Formula representing covariates on the initial state occupancy probabilities, via multinomial logistic regression. The linear effects of these covariates, observed at the individual's first observation time, operate on the log ratio of the state $r$ occupancy probability to the state 1 occupancy probability, for each $r=2$ to the number of states. Thus the state 1 occupancy probability should be non-zero. If est.initprobs is TRUE, these effects are estimated starting from their initial values. If est. initprobs is FALSE, these effects are fixed at theit initial values.
initcovinits Initial values for the covariate effects initcovariates. A named list with each element corresponding to a covariate, as in covinits. Each element is a vector with ( 1 - number of states) elements, containing the initial values for the linear effect of that covariate on the log odds of that state relative to state 1 , from state 2 to the final state. If initcovinits is not specified, all covariate effects are initialised to zero.
deathexact Vector of indices of absorbing states whose time of entry is known exactly, but the individual is assumed to be in an unknown transient state ("alive") at the
previous instant. This is the usual situation for times of death in chronic disease monitoring data. For example, if you specify deathexact $=c(4,5)$ then states 4 and 5 are assumed to be exactly-observed death states.
See the obstype argument. States of this kind correspond to obstype=3. deathexact $=$ TRUE indicates that the final absorbing state is of this kind, and deathexact = FALSE or deathexact $=$ NULL (the default) indicates that there is no state of this kind.
The deathexact argument is overridden by obstype or exacttimes.
Note that you do not always supply a deathexact argument, even if there are states that correspond to deaths, because they do not necessarily have obstype=3. If the state is known between the time of death and the previous observation, then you should specify obstype=2 for the death times, or exacttimes=TRUE if the state is known at all times, and the deathexact argument is ignored.
death Old name for the deathexact argument. Overridden by deathexact if both are supplied. Deprecated.
censor A state, or vector of states, which indicates censoring. Censoring means that the observed state is known only to be one of a particular set of states. For example, censor=999 indicates that all observations of 999 in the vector of observed states are censored states. By default, this means that the true state could have been any of the transient (non-absorbing) states. To specify corresponding true states explicitly, use a censor. states argument.
Note that in contrast to the usual terminology of survival analysis, here it is the state which is considered to be censored, rather than the event time. If at the end of a study, an individual has not died, but their true state is known, then censor is unnecessary, since the standard multi-state model likelihood is applicable. Also a "censored" state here can be at any time, not just at the end.
Note in particular that general time-inhomogeneous Markov models with piecewise constant transition intensities can be constructed using the censor facility. If the true state is unknown on occasions when a piecewise constant covariate is known to change, then censored states can be inserted in the data on those occasions. The covariate may represent time itself, in which case the pci option to msm can be used to perform this trick automatically, or some other timedependent variable.
censor.states Specifies the underlying states which censored observations can represent. If censor is a single number (the default) this can be a vector, or a list with one element. If censor is a vector with more than one element, this should be a list, with each element a vector corresponding to the equivalent element of censor. For example
censor $=c(99,999)$, censor. states $=\operatorname{list}(c(2,3), c(3,4))$
means that observations coded 99 represent either state 2 or state 3 , while observations coded 999 are really either state 3 or state 4 .
pci Model for piecewise-constant intensities. Vector of cut points defining the times, since the start of the process, at which intensities change for all subjects. For example
pci $=c(5,10)$
specifies that the intensity changes at time points 5 and 10 . This will automatically construct a model with a categorical (factor) covariate called timeperiod,
with levels "[-Inf,5)", "[5,10)" and "[10, Inf)", where the first level is the baseline. This covariate defines the time period in which the observation was made. Initial values and constraints on covariate effects are specified the same way as for a model with a covariate of this name, for example,
covinits $=$ list ("timeperiod[5,10)"=c(0.1,0.1),"timeperiod[10, Inf)" $=c(0.1,0.1)$ )
Thus if pci is supplied, you cannot have a previously-existing variable called timeperiod as a covariate in any part of a msm model.
To assume piecewise constant intensities for some transitions but not others with pci, use the fixedpars argument to fix the appropriate covariate effects at their default initial values of zero.
Internally, this works by inserting censored observations in the data at times when the intensity changes but the state is not observed.
If the supplied times are outside the range of the time variable in the data, pci is ignored and a time-homogeneous model is fitted.
After fitting a time-inhomogeneous model, qmatrix.msm can be used to obtain the fitted intensity matrices for each time period, for example,
qmatrix.msm(example.msm, covariates=list(timeperiod="[5,Inf)"))
This facility does not support interactions between time and other covariates. Such models need to be specified "by hand", using a state variable with censored observations inserted. Note that the data component of the msm object returned from a call to msm with pci supplied contains the states with inserted censored observations and time period indicators. These can be used to construct such models.
Note that you do not need to use pci in order to model the effect of a timedependent covariate in the data. msm will automatically assume that covariates are piecewise-constant and change at the times when they are observed. pci is for when you want all intensities to change at the same pre-specified times for all subjects.
phase.states Indices of states which have a two-phase sojourn distribution. This defines a semi-Markov model, in which the hazard of an onward transition depends on the time spent in the state.
This uses the technique described by Titman and Sharples (2009). A hidden Markov model is automatically constructed on an expanded state space, where the phases correspond to the hidden states. The "tau" proportionality constraint described in this paper is currently not supported.
Covariates, constraints, deathexact and censor are expressed with respect to the expanded state space. If not supplied by hand, initprobs is defined automatically so that subjects are assumed to begin in the first of the two phases.
Hidden Markov models can additionally be given phased states. The user supplies an outcome distribution for each original state using hmodel, which is expanded internally so that it is assumed to be the same within each of the phased states. initprobs is interpreted on the expanded state space. Misclassification models defined using ematrix are not supported, and these must be defined using hmmCat or hmmIdent constructors, as described in the hmodel section of this help page. Or the HMM on the expanded state space can be defined by hand.
Output functions are presented as it were a hidden Markov model on the expanded state space, for example, transition probabilities between states, covari-
ate effects on transition rates, or prevalence counts, are not aggregated over the hidden phases.
Numerical estimation will be unstable when there is weak evidence for a twophase sojourn distribution, that is, if the model is close to Markov.
See d2phase for the definition of the two-phase distribution and the interpretation of its parameters.
This is an experimental feature, and some functions are not implemented. Please report any experiences of using this feature to the author!
phase.inits Initial values for phase-type models. A list with one component for each "twophased" state. Each component is itself a list of two elements. The first of these elements is a scalar defining the transition intensity from phase 1 to phase 2. The second element is a matrix, with one row for each potential destination state from the two-phased state, and two columns. The first column is the transition rate from phase 1 to the destination state, and the second column is the transition rate from phase 2 to the destination state. If there is only one destination state, then this may be supplied as a vector.
In phase type models, the initial values for transition rates out of non-phased states are taken from the qmatrix supplied to msm, and entries of this matrix corresponding to transitions out of phased states are ignored.
exacttimes By default, the transitions of the Markov process are assumed to take place at unknown occasions in between the observation times. If exacttimes is set to TRUE, then the observation times are assumed to represent the exact times of transition of the process. The subject is assumed to be in the same state between these times. An observation may represent a transition to a different state or a repeated observation of the same state (e.g. at the end of follow-up). This is equivalent to every row of the data having obstype $=2$. See the obstype argument. If both obstype and exacttimes are specified then exacttimes is ignored.
Note that the complete history of the multi-state process is known with this type of data. The models which msm fits have the strong assumption of constant (or piecewise-constant) transition rates. Knowing the exact transition times allows more realistic models to be fitted with other packages. For example parametric models with sojourn distributions more flexible than the exponential can be fitted with the flexsurv package, or semi-parametric models can be implemented with survival in conjunction with mstate.
cl Width of symmetric confidence intervals for maximum likelihood estimates, by default 0.95 .
fixedpars Vector of indices of parameters whose values will be fixed at their initial values during the optimisation. These are given in the order: transition intensities (reading across rows of the transition matrix), covariates on intensities (ordered by intensities within covariates), hidden Markov model parameters, including misclassification probabilities or parameters of HMM outcome distributions (ordered by parameters within states), hidden Markov model covariate parameters (ordered by covariates within parameters within states), initial state occupancy probabilities (excluding the first probability, which is fixed at one minus the sum of the others).
If there are equality constraints on certain parameters, then fixedpars indexes
the set of unique parameters, excluding those which are constrained to be equal
to previous parameters.
To fix all parameters, specify fixedpars = TRUE.
This can be useful for profiling likelihoods, and building complex models stage
by stage.
If TRUE (the default, unless fixedpars=TRUE) then covariates are centered at
their means during the maximum likelihood estimation. This usually improves
stability of the numerical optimisation.
opt.method
If "optim", "nlm" or "bobyqa", then the corresponding R function will be used
for maximum likelihood estimation. optim is the default. "bobyqa" requires
the package minqa to be installed. See the help of these functions for further
details. Advanced users can also add their own optimisation methods, see the
source for optim.R in msm for some examples.
If "fisher", then a specialised Fisher scoring method is used (Kalbfleisch and
Lawless, 1985) which can be faster than the generic methods, though less robust.
This is only available for Markov models with panel data (obstype=1), that is,
nessian
not for models with censored states, hidden Markov models, exact observation
or exact death times (obstype=2,3).
If TRUE then standard errors and confidence intervals are obtained from a nu-
merical estimate of the Hessian (the observed information matrix). This is the
default when maximum likelihood estimation is performed. If all parameters are
fixed at their initial values and no optimisation is performed, then this defaults to
all NAs in other covariates (excluding the last observation for a subject), all NAs in obstype (excluding the first observation for a subject), and any subjects with only one observation (thus no observed transitions).
Optional arguments to the general-purpose $R$ optimisation routine, optim by default. For example method="Nelder-Mead" to change the optimisation algorithm from the "BFGS" method that msm calls by default.
It is often worthwhile to normalize the optimisation using control=list (fnscale $=a$ ), where $a$ is the a number of the order of magnitude of the -2 log likelihood.
If 'false' convergence is reported and the standard errors cannot be calculated due to a non-positive-definite Hessian, then consider tightening the tolerance criteria for convergence. If the optimisation takes a long time, intermediate steps can be printed using the trace argument of the control list. See optim for details.
For the Fisher scoring method, a control list can be supplied in the same way, but the only supported options are reltol, trace and damp. The first two are used in the same way as for optim. If the algorithm fails with a singular information matrix, adjust damp from the default of zero (to, e.g. 1). This adds a constant identity matrix multiplied by damp to the information matrix during optimisation.

## Details

For full details about the methodology behind the msm package, refer to the PDF manual 'msm-manual . pdf' in the 'doc' subdirectory of the package. This includes a tutorial in the typical use of $\mathbf{m s m}$. The paper by Jackson (2011) in Journal of Statistical Software presents the material in this manual in a more concise form.
msm was designed for fitting continuous-time Markov models, processes where transitions can occur at any time. These models are defined by intensities, which govern both the time spent in the current state and the probabilities of the next state. In discrete-time models, transitions are known in advance to only occur at multiples of some time unit, and the model is purely governed by the probability distributions of the state at the next time point, conditionally on the state at the current time. These can also be fitted in msm, assuming that there is a continuous-time process underlying the data. Then the fitted transition probability matrix over one time period, as returned by pmatrix.msm(..,$t=1$ ) is equivalent to the matrix that governs the discrete-time model. However, these can be fitted more efficiently using multinomial logistic regression, for example, using multinom from the R package nnet (Venables and Ripley, 2002).

For simple continuous-time multi-state Markov models, the likelihood is calculated in terms of the transition intensity matrix $Q$. When the data consist of observations of the Markov process at arbitrary times, the exact transition times are not known. Then the likelihood is calculated using the transition probability matrix $P(t)=\exp (t Q)$, where $\exp$ is the matrix exponential. If state $i$ is observed at time $t$ and state $j$ is observed at time $u$, then the contribution to the likelihood from this pair of observations is the $i, j$ element of $P(u-t)$. See, for example, Kalbfleisch and Lawless (1985), Kay (1986), or Gentleman et al. (1994).

For hidden Markov models, the likelihood for an individual with $k$ observations is calculated directly by summing over the unknown state at each time, producing a product of $k$ matrices. The calculation is a generalisation of the method described by Satten and Longini (1996), and also by Jackson and Sharples (2002), and Jackson et al. (2003).

There must be enough information in the data on each state to estimate each transition rate, otherwise the likelihood will be flat and the maximum will not be found. It may be appropriate to reduce the number of states in the model, the number of allowed transitions, or the number of covariate effects, to ensure convergence. Hidden Markov models, and situations where the value of the process is only known at a series of snapshots, are particularly susceptible to non-identifiability, especially when combined with a complex transition matrix. Choosing an appropriate set of initial values for the optimisation can also be important. For flat likelihoods, 'informative' initial values will often be required. See the PDF manual for other tips.

## Value

To obtain summary information from models fitted by the msm function, it is recommended to use extractor functions such as qmatrix.msm, pmatrix.msm, sojourn.msm, msm.form.qoutput. These provide estimates and confidence intervals for quantities such as transition probabilities for given covariate values.

For advanced use, it may be necessary to directly use information stored in the object returned by msm. This is documented in the help page msm. object.
Printing a msm object by typing the object's name at the command line implicitly invokes print.msm. This formats and prints the important information in the model fit, and also returns that information in an R object. This includes estimates and confidence intervals for the transition intensities and (log) hazard ratios for the corresponding covariates. When there is a hidden Markov model, the chief information in the hmodel component is also formatted and printed. This includes estimates and confidence intervals for each parameter.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## References

Jackson, C.H. (2011). Multi-State Models for Panel Data: The msm Package for R., Journal of Statistical Software, 38(8), 1-29. URL http://www.jstatsoft.org/v38/i08/.

Kalbfleisch, J., Lawless, J.F., The analysis of panel data under a Markov assumption Journal of the Americal Statistical Association (1985) 80(392): 863-871.
Kay, R. A Markov model for analysing cancer markers and disease states in survival studies. Biometrics (1986) 42: 855-865.

Gentleman, R.C., Lawless, J.F., Lindsey, J.C. and Yan, P. Multi-state Markov models for analysing incomplete disease history data with illustrations for HIV disease. Statistics in Medicine (1994) 13(3): 805-821.

Satten, G.A. and Longini, I.M. Markov chains with measurement error: estimating the 'true' course of a marker of the progression of human immunodeficiency virus disease (with discussion) Applied Statistics 45(3): 275-309 (1996)
Jackson, C.H. and Sharples, L.D. Hidden Markov models for the onset and progression of bronchiolitis obliterans syndrome in lung transplant recipients Statistics in Medicine, 21(1): 113-128 (2002).

Jackson, C.H., Sharples, L.D., Thompson, S.G. and Duffy, S.W. and Couto, E. Multi-state Markov models for disease progression with classification error. The Statistician, 52(2): 193-209 (2003)

Titman, A.C. and Sharples, L.D. Semi-Markov models with phase-type sojourn distributions. Biometrics 66, 742-752 (2009).
Venables, W.N. and Ripley, B.D. (2002) Modern Applied Statistics with S, second edition. Springer.

## See Also

simmulti.msm, plot.msm, summary.msm, qmatrix.msm, pmatrix.msm, sojourn.msm.

## Examples

```
### Heart transplant data
### For further details and background to this example, see
### Jackson (2011) or the PDF manual in the doc directory.
print(cav[1:10,])
twoway4.q <- rbind(c(-0.5, 0.25, 0, 0.25), c(0.166, -0.498, 0.166, 0.166),
c(0, 0.25, -0.5, 0.25), c(0, 0, 0, 0))
statetable.msm(state, PTNUM, data=cav)
crudeinits.msm(state ~ years, PTNUM, data=cav, qmatrix=twoway4.q)
cav.msm <- msm( state ~ years, subject=PTNUM, data = cav,
                        qmatrix = twoway4.q, deathexact = 4,
        control = list ( trace = 2, REPORT = 1 ) )
cav.msm
qmatrix.msm(cav.msm)
pmatrix.msm(cav.msm, t=10)
sojourn.msm(cav.msm)
```

msm.form.qoutput Extract msm model parameter estimates in compact format

## Description

Extract estimates and confidence intervals for transition intensities (or misclassification probabilities), and their covariate effects, in a tidy matrix format with one row per transition. This is used by the print method (print.msm) for msm objects. Covariate effects are returned as hazard or odds ratios, not on the $\log$ scale.

## Usage

msm.form.qoutput(x, covariates="mean", cl=0.95, digits=4, ...)
msm.form.eoutput(x, covariates="mean", cl=0.95, digits=4, ...)

## Arguments

x
A fitted multi-state model object, as returned by msm.
covariates
cl
digits Covariate values defining the "baseline" parameters (see qmatrix.msm). Width of the symmetric confidence interval to present. Defaults to 0.95 .
Minimum number of significant digits for the formatted character matrix re- turned as an attribute. This is passed to format. Defaults to 4.
... Other arguments to be passed to format.

## Value

A numeric matrix with one row per transition, and one column for each estimate or confidence limit. The "formatted" attribute contains the same results formatted for pretty printing. msm. form. qoutput returns the transition intensities and their covariates, and msm. form. eoutput returns the misclassification probabilities and their covariates.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

See Also<br>print.msm

msm.object Fitted msm model objects

## Description

The msm function returns a list with the following components. These are intended for developers and confident users. To extract results from fitted model objects, functions such as qmatrix.msm or print.msm should be used instead.

## Value

call The original call to msm, as returned by match.call.
Qmatrices A list of matrices. The first component, labelled logbaseline, is a matrix containing the estimated transition intensities on the log scale with any covariates fixed at their means in the data (or at zero, if center=FALSE). The component labelled baseline is the equivalent on the untransformed scale. Each remaining component is a matrix giving the linear effects of the labelled covariate on the matrix of log intensities. To extract an estimated intensity matrix on the natural scale, at an arbitrary combination of covariate values, use the function qmatrix.msm.
QmatricesSE The standard error matrices corresponding to Qmatrices.
QmatricesL, QmatricesU
Corresponding lower and upper symmetric confidence limits, of width 0.95 unless specified otherwise by the cl argument.
Ematrices A list of matrices. The first component, labelled logitbaseline, is the estimated misclassification probability matrix (expressed as as log odds relative to the probability of the true state) with any covariates fixed at their means in the data (or at zero, if center=FALSE). The component labelled baseline is the equivalent on the untransformed scale. Each remaining component is a matrix giving the linear effects of the labelled covariate on the matrix of logit misclassification probabilities. To extract an estimated misclassification probability matrix on the natural scale, at an arbitrary combination of covariate values, use the function ematrix.msm.
\(\left.\begin{array}{ll}EmatricesSE \& The standard error matrices corresponding to Ematrices. <br>
EmatricesL, EmatricesU <br>
Corresponding lower and upper symmetric confidence limits, of width 0.95 un- <br>

less specified otherwise by the cl argument.\end{array}\right]\)| Minus twice the maximised log-likelihood. |
| :--- |
| minus2loglik |
| deriv |
| estimates |
| Derivatives of the minus twice log-likelihood at its maximum. |
| Vector of untransformed maximum likelihood estimates returned from optim. |
| Transition intensities are on the log scale and misclassification probabilities are |
| given as log odds relative to the probability of the true state. |
| Vector of transformed maximum likelihood estimates with intensities and prob- |
| abilities on their natural scales. |

paramdata A list giving information about the parameters of the multi-state model. See paramdata.object.
cl Confidence interval width, as supplied to msm.
covariates Formula for covariates on intensities, as supplied to msm.
misccovariates Formula for covariates on misclassification probabilities, as supplied to msm.
hcovariates Formula for covariates on hidden Markov model outcomes, as supplied to msm.
initcovariates Formula for covariates on initial state occupancy probabilities in hidden Markov models, as supplied to msm.
sojourn A list as returned by sojourn.msm, with components:
mean = estimated mean sojourn times in the transient states, with covariates fixed at their means (if center=TRUE) or at zero (if center=FALSE).
se $=$ corresponding standard errors.

```
msm.summary
```

Summarise a fitted multi-state model

## Description

Summary method for fitted msm models. This is simply a wrapper around prevalence.msm which produces a table of observed and expected state prevalences for each time, and for models with covariates, hazard.msm to print hazard ratios with $95 \%$ confidence intervals for covariate effects.

## Usage

\#\# S3 method for class 'msm'
summary (object, hazard.scale=1, ...)

## Arguments

object A fitted multi-state model object, as returned by msm.
hazard.scale Vector with same elements as number of covariates on transition rates. Corresponds to the increase in each covariate used to calculate its hazard ratio. Defaults to all 1 .
... Further arguments passed to prevalence.msm.

## Value

A list of class summary.msm, with components:
prevalences Output from prevalence.msm.
hazard Output from hazard.msm.
hazard.scale Value of the hazard.scale argument.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

msm,prevalence.msm, hazard.msm

$$
\begin{array}{ll}
\text { msm2Surv } & \begin{array}{l}
\text { Convert data for 'msm' to data for 'survival', 'mstate' or 'flexsurv' } \\
\text { analysis }
\end{array}
\end{array}
$$

## Description

Converts longitudinal data for a msm model fit, where observations represent the exact transition times of the process, to counting process data. This enables, for example, flexible parametric multistate models to be fitted with flexsurvreg from the flexsurv package, or semiparametric models to be implemented with coxph and the mstate package.

## Usage

msm2Surv(data, subject, time, state, covs, Q)

## Arguments

data Data frame in the format expected by a msm model fit with exacttimes=TRUE or all obstype=2. Each row represents an observation of a state, and the time variable contains the exact and complete transition times of the underlying process. This is explained in more detail in the help page for msm, section obstype=2.
subject Name of the subject ID in the data (character format, i.e. quoted).
time Name of the time variable in the data (character).
state $\quad$ Name of the state variable in the data (character).
covs Vector of covariate names to carry through (character). If not supplied, this is taken to be all remaining variables in the data.
Q Transition intensity matrix. This should have number of rows and number of columns both equal to the number of states. If an instantaneous transition is not allowed from state $r$ to state $s$, then Q should have $(r, s)$ entry 0 , otherwise it should be non-zero. The diagonal entries are ignored.

## Details

For example, if the data supplied to msm look like this:

| subj | days | status | age | treat |
| :--- | :--- | :--- | :--- | :--- |
| 1 | 0 | 1 | 66 | 1 |
| 1 | 27 | 2 | 66 | 1 |
| 1 | 75 | 3 | 66 | 1 |


| 1 | 97 | 4 | 66 | 1 |
| :--- | :--- | :--- | :--- | :--- |
| 1 | 1106 | 4 | 69 | 1 |
| 2 | 0 | 1 | 49 | 0 |
| 2 | 90 | 2 | 49 | 0 |
| 2 | 1037 | 2 | 51 | 0 |

then the output of msm2Surv will be a data frame looking like this:

| id | from | to | Tstart | Tstop | time | status | age | treat | trans |
| :---: | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 1 | 2 | 0 | 27 | 27 | 1 | 66 | 1 | 1 |
| 1 | 1 | 4 | 0 | 27 | 27 | 0 | 66 | 1 | 2 |
| 1 | 2 | 3 | 27 | 75 | 48 | 1 | 66 | 1 | 3 |
| 1 | 2 | 4 | 27 | 75 | 48 | 0 | 66 | 1 | 4 |
| 1 | 3 | 4 | 75 | 97 | 22 | 1 | 69 | 1 | 5 |
| 2 | 1 | 2 | 0 | 90 | 90 | 1 | 49 | 0 | 1 |
| 2 | 1 | 4 | 0 | 90 | 90 | 0 | 49 | 0 | 2 |
| 2 | 2 | 3 | 90 | 1037 | 947 | 0 | 49 | 0 | 3 |
| 2 | 2 | 4 | 90 | 1037 | 947 | 0 | 49 | 0 | 4 |

At 27 days, subject 1 is observed to move from state 1 to state 2 (first row, status 1 ), which means that their potential transition from state 1 to state 4 is censored (second row, status 0 ).
See the mstate package and the references below for more details of this data format and using it for semi-parametric multi-state modelling.

## Value

A data frame of class "msdata", with rows representing observed or censored transitions. There will be one row for each observed transition in the original data, and additional rows for every potential transition that could have occurred out of each observed state.
The data frame will have columns called:

| id | Subject ID |
| :---: | :---: |
| from | Starting state of the transition |
| to | Finishing state of the transition |
| Tstart | The starting time of the transition |
| Tstop | The finishing time of the transition |
| time | The time difference = Tstop - Tstart |
| status | Event or censoring indicator, with 1 indicating an observed transition, and 0 indicating censoring |
| trans | Transition number |
| and any the stand "status | g columns will represent covariates. Any covariates whose names clash with bles in the returned data ("id", "from", "to", "Tstart", "Tstop", "time", ns") have ". 2 " appended to their names. |

The transition matrix in mstate format is stored in the trans attribute of the returned object. See the example code below.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## References

Putter H, Fiocco M, Geskus RB (2007). Tutorial in biostatistics: Competing risks and multi-state models. Statistics in Medicine 26: 2389-2430.

Liesbeth C. de Wreede, Marta Fiocco, Hein Putter (2011). mstate: An R Package for the Analysis of Competing Risks and Multi-State Models. Journal of Statistical Software, 38(7), 1-30. https: //www.jstatsoft.org/v38/i07

Jackson, C. H. (2014). flexsurv: Flexible parametric survival and multi-state models. R package version 0.5.

## See Also

msprep, in mstate, which produces data in a similar format, given data in "wide" format with one row per subject.

## Examples

```
msmdat <- data.frame(
    subj = c(1, 1, 1, 1, 1, 2, 2, 2),
    days = c(0, 27, 75, 97, 1106, 0, 90, 1037),
    status = c(1, 2, 3, 4, 4, 1, 2, 2),
    age =c(66, 66, 66, 66, 69, 49, 49, 51),
    treat = c(1, 1, 1, 1, 1, 0, 0, 0)
)
# transitions only allowed to next state up or state 4
Q <- rbind(c(1, 1, 0, 1),
            c(0, 1, 1, 1),
            c(0, 0, 1, 1),
            c(0, 0, 0, 0))
dat <- msm2Surv(data=msmdat, subject="subj", time="days", state="status",
            Q=Q)
dat
attr(dat, "trans")
```

odds.msm Calculate tables of odds ratios for covariates on misclassification probabilities

## Description

Odds ratios are computed by exponentiating the estimated covariate effects on the logit-misclassification probabilities.

## Usage

odds.msm(x, odds.scale = 1, cl = 0.95)

## Arguments

x
odds.scale Vector with same elements as number of covariates on misclassification probabilities. Corresponds to the increase in each covariate used to calculate its odds ratio. Defaults to all 1.
cl Width of the symmetric confidence interval to present. Defaults to 0.95 .

## Value

A list of tables containing odds ratio estimates, one table for each covariate. Each table has three columns, containing the odds ratio, and an approximate upper $95 \%$ and lower $95 \%$ confidence limit respectively (assuming normality on the log scale), for each misclassification probability.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

msm, hazard.msm
paramdata.object Developer documentation: internal msm parameters object

## Description

An object giving information about the parameters of the multi-state model. Used internally during maximum likelihood estimation and arranging results. Returned as the paramdata component of a fitted msm model object.

## Value

| inits | Vector of initial values for distinct parameters which are being estimated. These <br> have been transformed to the real line (e.g. by log), and exclude parameters <br> being fixed at their initial values, parameters defined to be always fixed (e.g. <br> binomial denominators) and parameters constrained to equal previous ones. |
| :--- | :--- |
| plabs | Names of parameters in allinits. |
| allinits | Vector of parameter values before estimation, including those which are fixed or <br> constrained to equal other parameters, and transformed to the real line. |
| hmmpars | Indices of allinits which represent baseline parameters of hidden Markov out- <br> come models (thus excluding covariate effects in HMMs and initial state occu- <br> pancy probabilities). |


| fixed | TRUE if all parameters are fixed, FALSE otherwise. |
| :---: | :---: |
| fixedpars | Indices of parameters in allinits which are fixed, either by definition or as requested by the user in the fixedpars argument to msm. Excludes parameters fixed by constraining to equal other parameters. |
| notfixed | Indices of parameters which are not fixed by the definition of fixedpars. |
| optpars | Indices of parameters in allinits being estimated, thus those included in inits. |
| auxpars | Indices of "auxiliary" parameters which are always fixed, for example, binomial denominators (hmmBinom) and the which parameter in hmmIdent. |
| constr | Vector of integers, of length npars, indicating which sets of parameters are constrained to be equal to each other. If two of these integers are equal the corresponding parameters are equal. A negative element indicates that parameter is defined to be minus some other parameter (this is used for covariate effects on transition intensities). |
| npars | Total number of parameters, equal to length(allinits). |
| nfix | Number of fixed parameters, equal to length(fixedpars). |
| nopt | Number of parameters being estimated, equal to length(inits) and length(optpars). |
| ndup | Number of parameters defined as duplicates of previous parameters by equality constraints (currently unused). |
| ranges | Matrix of defined ranges for each parameter on the natural scale (e.g. 0 to infinity for rate parameters). |
| opt | Object returned by the optimisation routine (such as optim). |
| foundse | TRUE if standard errors are available after optimisation. If FALSE the optimisation probably hasn't converged. |
| lik | Minus twice the log likelihood at the parameter estimates. |
| deriv | Derivatives of the minus twice log likelihood at the parameter estimates, if available. |
| information | Corresponding expected information matrix at the parameter estimates, if available. |
| params | Vector of parameter values after maximum likelihood estimation, corresponding to allinits, still on the real-line transformed scale. |
| covmat | Covariance matrix corresponding to params. |
| ci | Matrix of confidence intervals corresponding to params, with nominal coverage (default 0.95 ) defined by the cl argument of msm. |
| estimates.t | Vector of parameter estimates, as params but with parameters on their natural scales. |

## See Also

msm.object

## Description

Pearson-type goodness-of-fit test for multi-state models fitted to panel-observed data.

## Usage

pearson.msm(x, transitions=NULL, timegroups=3, intervalgroups=3, covgroups=3, groups=NULL, boot=FALSE, B=500, next.obstime=NULL, $N=100$, indep.cens=TRUE, maxtimes=NULL, pval=TRUE)

## Arguments

$x \quad$ A fitted multi-state model, as returned by msm.
transitions This should be an integer vector indicating which interval transitions should be grouped together in the contingency table. Its length should be the number of allowed interval transitions, excluding transitions from absorbing states to absorbing states.
The allowed interval transitions are the set of pairs of states $(a, b)$ for which it is possible to observe $a$ at one time and $b$ at any later time. For example, in a "well-disease-death" model with allowed instantaneous 1-2, 2-3 transitions, there are 5 allowed interval transitions. In numerical order, these are 1-1, 1-2, 1-3, 2-2 and 2-3, excluding absorbing-absorbing transitions.
Then, to group transitions 1-1,1-2 together, and transitions 2-2,2-3 together, specify
transitions $=c(1,1,2,3,3)$.
Only transitions from the same state may be grouped. By default, each interval transition forms a separate group.
timegroups Number of groups based on quantiles of the time since the start of the process.
intervalgroups Number of groups based on quantiles of the time interval between observations, within time groups
covgroups $\quad$ Number of groups based on quantiles of $\sum_{r} q_{i r r}$, where $q_{i r r}$ are the diagonal entries of the transition intensity matrix for the $i$ th transition. These are a function of the covariate effects and the covariate values at the $i$ th transition: $q_{i r r}$ is minus the sum of the off-diagonal entries $q_{r s}^{(0)} \exp \left(\beta_{r s}^{T} z_{i}\right)$ on the $r$ th row.
Thus covgroups summarises the impact of covariates at each observation, by calculating the overall rate of progression through states at that observation.
For time-inhomogeneous models specified using the pci argument to msm, if the only covariate is the time period, covgroups is set to 1 , since timegroups ensures that transitions are grouped by time.
groups A vector of arbitrary groups in which to categorise each transition. This can be an integer vector or a factor. This can be used to diagnose specific areas of poor fit. For example, the contingency table might be grouped by arbitrary combinations of covariates to detect types of individual for whom the model fits poorly.
The length of groups should be $\mathrm{x} \$$ data\$n, the number of observations used in the model fit, which is the number of observations in the original dataset with any missing values excluded. The value of groups at observation $i$ is used to categorise the transition which ends at observation i. Values of groups at the first observation for each subject are ignored.
boot Estimate an "exact" p-value using a parametric bootstrap.
All objects used in the original call to msm which produced $x$, such as the qmatrix, should be in the working environment, or else an "object not found" error will be given. This enables the original model to be refitted to the replicate datasets. Note that groups cannot be used with bootstrapping, as the simulated observations will not be in the same categories as the original observations.
B
Number of bootstrap replicates.
next.obstime This is a vector of length $x \$ d a t a \$ n$ (the number of observations used in the model fit) giving the time to the next scheduled observation following each time point. This is only used when times to death are known exactly.
For individuals who died (entered an absorbing state) before the next scheduled observation, and the time of death is known exactly, next. obstime would be greater than the observed death time.
If the individual did not die, and a scheduled observation did follow that time point, next. obstime should just be the same as the time to that observation.
next.obstime is used to determine a grouping of the time interval between observations, which should be based on scheduled observations. If exact times to death were used in the grouping, then shorter intervals would contain excess deaths, and the goodness-of-fit statistic would be biased.
If next. obstime is unknown, it is multiply-imputed using a product-limit estimate based on the intervals to observations other than deaths. The resulting tables of transitions are averaged over these imputations. This may be slow.
N
indep.cens If TRUE, then times to censoring are included in the estimation of the distribution to the next scheduled observation time. If FALSE, times to censoring are assumed to be systematically different from other observation times.
maxtimes A vector of length $\times \$$ data $\$ n$, or a common scalar, giving an upper bound for the next scheduled observation time. Used in the multiple imputation when times to death are known exactly. If a value greater than maxtimes is simulated, then the next scheduled observation is taken as censored. This should be supplied, if known. If not supplied, this is taken to be the maximum interval occurring in the data, plus one time unit. For observations which are not exact death times, this should be the time since the previous observation.
pval Calculate a p-value using the improved approximation of Titman (2009). This is optional since it is not needed during bootstrapping, and it is computationally
non-trivial. Only available currently for non-hidden Markov models for panel data without exact death times. Also not available for models with censoring, including time-homogeneous models fitted with the pci option to msm.

## Details

This method (Aguirre-Hernandez and Farewell, 2002) is intended for data which represent observations of the process at arbitrary times ("snapshots", or "panel-observed" data). For data which represent the exact transition times of the process, prevalence.msm can be used to assess fit, though without a formal test.

When times of death are known exactly, states are misclassified, or an individual's final observation is a censored state, the modification by Titman and Sharples (2008) is used. The only form of censoring supported is a state at the end of an individual's series which represents an unknown transient state (i.e. the individual is only known to be alive at this time). Other types of censoring are omitted from the data before performing the test.
See the references for further details of the methods. The method used for censored states is a modification of the method in the appendix to Titman and Sharples (2008), described at http: //www.mrc-bsu.cam.ac.uk/wp-content/uploads/robustcensoring.pdf (Titman, 2007).

Groupings of the time since initiation, the time interval and the impact of covariates are based on equally-spaced quantiles. The number of groups should be chosen that there are not many cells with small expected numbers of transitions, since the deviance statistic will be unstable for sparse contingency tables. Ideally, the expected numbers of transitions in each cell of the table should be no less than about 5 . Conversely, the power of the test is reduced if there are too few groups. Therefore, some sensitivity analysis of the test results to the grouping is advisable.

Saved model objects fitted with previous versions of R (versions less than 1.2 ) will need to be refitted under the current R for use with pearson.msm.

## Value

A list whose first two elements are contingency tables of observed transitions $O$ and expected transitions $E$, respectively, for each combination of groups. The third element is a table of the deviances $(O-E)^{2} / E$ multiplied by the sign of $O-E$. If the expected number of transitions is zero then the deviance is zero. Entries in the third matrix will be bigger in magnitude for groups for which the model fits poorly.
"test" the fourth element of the list, is a data frame with one row containing the Pearson-type goodness-of-fit test statistic stat. The test statistic is the sum of the deviances. For panel-observed data without exact death times, misclassification or censored observations, $p$ is the p-value for the test statistic calculated using the improved approximation of Titman (2009).
For these models, for comparison with older versions of the package, test also presents p.lower and p.upper, which are theoretical lower and upper limits for the p-value of the test statistic, based on $\chi^{2}$ distributions with df . lower and df . upper degrees of freedom, respectively. df. upper is the number of independent cells in the contingency table, and df. lower is df. upper minus the number of estimated parameters in the model.

```
"intervalq" (not printed by default) contains the definition of the grouping of the intervals
        between observations. These groups are defined by quantiles within the groups
        corresponding to the time since the start of the process.
"sim" If there are exact death times, this contains simulations of the contingency ta-
        bles and test statistics for each imputation of the next scheduled sampling time.
        These are averaged over to produce the presented tables and test statistic. This
        element is not printed by default.
        With exact death times, the null variance of the test statistic (formed by taking
        mean of simulated test statistics) is less than twice the mean (Titman, 2008), and
        the null distribution is not \(\chi^{2}\). In this case, \(p\). upper is an upper limit for the true
        asymptotic p-value, but p. lower is not a lower limit, and is not presented.
"boot" If the bootstrap has been used, the element will contain the bootstrap replicates
        of the test statistics (not printed by default).
"lambda" If the Titman (2009) p-value has been calculated, this contains the weights defin-
        ing the null distribution of the test statistic as a weighted sum of \(\chi_{1}^{2}\) random
        variables (not printed by default).
```


## Author(s)

Andrew Titman [a.titman@lancaster.ac.uk](mailto:a.titman@lancaster.ac.uk), Chris Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## References

Aguirre-Hernandez, R. and Farewell, V. (2002) A Pearson-type goodness-of-fit test for stationary and time-continuous Markov regression models. Statistics in Medicine 21:1899-1911.
Titman, A. and Sharples, L. (2008) A general goodness-of-fit test for Markov and hidden Markov models. Statistics in Medicine 27(12):2177-2195
Titman, A. (2009) Computation of the asymptotic null distribution of goodness-of-fit tests for multistate models. Lifetime Data Analysis 15(4):519-533.
Titman, A. (2008) Model diagnostics in multi-state models of biological systems. PhD thesis, University of Cambridge.

## See Also

msm, prevalence.msm, scoreresid.msm,

## Examples

```
psor.q <- rbind(c(0,0.1,0,0),c(0,0,0.1,0),c(0,0,0,0.1),c(0,0,0,0))
psor.msm <- msm(state ~ months, subject=ptnum, data=psor,
    qmatrix = psor.q, covariates = ~ollwsdrt+hieffusn,
    constraint = list(hieffusn=c(1,1,1),ollwsdrt=c(1,1,2)))
pearson.msm(psor.msm, timegroups=2, intervalgroups=2, covgroups=2)
# More 1-2, 1-3 and 1-4 observations than expected in shorter time
# intervals - the model fits poorly.
# A random effects model might accommodate such fast progressors.
```


## Description

Density, distribution function, quantile function and random generation for a generalisation of the exponential distribution, in which the rate changes at a series of times.

## Usage

```
dpexp(x, rate=1, t=0, log = FALSE)
ppexp(q, rate=1, t=0, lower.tail = TRUE, log.p = FALSE)
qpexp(p, rate=1, t=0, lower.tail = TRUE, log.p = FALSE)
rpexp(n, rate=1, t=0)
```


## Arguments

$x, q \quad$ vector of quantiles.
$p \quad$ vector of probabilities.
$n \quad$ number of observations. If length $(n)>1$, the length is taken to be the number required.
rate vector of rates.
t
vector of the same length as rate, giving the times at which the rate changes. The first element of $t$ should be 0 , and $t$ should be in increasing order.
$\log , \log . \mathrm{p} \quad \operatorname{logical}$; if TRUE, probabilities p are given as $\log (\mathrm{p})$, or $\log$ density is returned.
lower.tail logical; if TRUE (default), probabilities are $\mathrm{P}[\mathrm{X}<=\mathrm{x}]$, otherwise, $\mathrm{P}[\mathrm{X}>\mathrm{x}]$.

## Details

Consider the exponential distribution with rates $r_{1}, \ldots, r_{n}$ changing at times $t_{1}, \ldots, t_{n}$, with $t_{1}=$ 0 . Suppose $t_{k}$ is the maximum $t_{i}$ such that $t_{i}<x$. The density of this distribution at $x>0$ is $f(x)$ for $k=1$, and

$$
\prod_{i=1}^{k}\left(1-F\left(t_{i}-t_{i-1}, r_{i}\right)\right) f\left(x-t_{k}, r_{k}\right)
$$

for $\mathrm{k}>1$.
where $F()$ and $f()$ are the distribution and density functions of the standard exponential distribution.
If rate is of length 1 , this is just the standard exponential distribution. Therefore, for example, $\operatorname{dpexp}(x)$, with no other arguments, is simply equivalent to $\operatorname{dexp}(x)$.

Only rpexp is used in the msm package, to simulate from Markov processes with piecewise-constant intensities depending on time-dependent covariates. These functions are merely provided for completion, and are not optimized for numerical stability or speed.

## Value

dpexp gives the density, ppexp gives the distribution function, qpexp gives the quantile function, and rpexp generates random deviates.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

dexp, sim.msm.

## Examples

```
x <- seq(0.1, 50, by=0.1)
rate <- c(0.1, 0.2, 0.05, 0.3)
t <- c(0, 10, 20, 30)
## standard exponential distribution
plot(x, dexp(x, 0.1), type="l")
## distribution with piecewise constant rate
lines(x, dpexp(x, rate, t), type="l", lty=2)
## standard exponential distribution
plot(x, pexp(x, 0.1), type="l")
## distribution with piecewise constant rate
lines(x, ppexp(x, rate, t), type="l", lty=2)
```

phasemeans.msm Parameters of phase-type models in mixture form

## Description

Parameters of fitted two-phase models, in mixture model parameterisation.

## Usage

phasemeans.msm(x, covariates="mean", ci=c("none","normal","bootstrap"), $\mathrm{cl}=0.95, \mathrm{~B}=1000$, cores=NULL)

## Arguments

x
A fitted multi-state model, as returned by msm.
covariates
ci
cl Width of the symmetric confidence interval, relative to 1.
B Number of bootstrap replicates, or number of normal simulations from the distribution of the MLEs.
cores Number of cores to use for bootstrapping using parallel processing. See boot.msm for more details.

## Value

Matrix with one row for each state that has a two-phase distribution, and three columns: the shortstay mean, long-stay mean and long-stay probability. These are functions of the transition intensities of the expanded hidden Markov model, defined in d2phase.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

d2phase.

```
plot.msm Plots of multi-state models
```


## Description

This produces a plot of the expected probability of survival against time, from each transient state. Survival is defined as not entering an absorbing state.

## Usage

```
## S3 method for class 'msm'
```

plot(x, from, to, range, covariates, legend.pos, xlab="Time",
ylab="Fitted survival probability", lwd=1, ...)

## Arguments

$x \quad$ Output from msm, representing a fitted multi-state model object.
from States from which to consider survival. Defaults to the complete set of transient states.
to Absorbing state to consider. Defaults to the highest-labelled absorbing state.
range Vector of two elements, giving the range of times to plot for.
covariates Covariate values for which to evaluate the expected probabilities. This can either be:
the string "mean", denoting the means of the covariates in the data (this is the default),
the number 0 , indicating that all the covariates should be set to zero,
or a list of values, with optional names. For example
list $(60,1)$
where the order of the list follows the order of the covariates originally given in the model formula, or a named list,
list (age $=60$, sex $=1$ )
legend.pos Vector of the $x$ and $y$ position, respectively, of the legend.
xlab $\quad x$ axis label.
ylab y axis label.
lwd Line width. See par.
... Other arguments to be passed to the generic plot and lines functions.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

msm
plot.prevalence.msm Plot of observed and expected prevalences

## Description

Provides a rough indication of goodness of fit of a multi-state model, by estimating the observed numbers of individuals occupying a state at a series of times, and plotting these against forecasts from the fitted model, for each state. Observed prevalences are indicated as solid lines, expected prevalences as dashed lines.

## Usage

```
## S3 method for class 'prevalence.msm'
plot(x, mintime=NULL, maxtime=NULL, timezero=NULL,
                        initstates=NULL, interp=c("start","midpoint"),
        censtime=Inf, subset=NULL,
        covariates="population", misccovariates="mean",
        piecewise.times=NULL, piecewise.covariates=NULL,
        xlab="Times",ylab="Prevalence (%)", lwd.obs=1,
        lwd.exp=1, lty.obs=1, lty.exp=2, col.obs="blue",
        col.exp="red", legend.pos=NULL,
        ...)
```


## Arguments

| x | A fitted multi-state model produced by msm. |
| :--- | :--- |
| mintime | Minimum time at which to compute the observed and expected prevalences of <br> states. |
| maxtime | Maximum time at which to compute the observed and expected prevalences of <br> states. |
| timezero | Initial time of the Markov process. Expected values are forecasted from here. <br> Defaults to the minimum of the observation times given in the data. |


| initstates | Optional vector of the same length as the number of states. Gives the numbers <br> of individuals occupying each state at the initial time, to be used for forecasting <br> expected prevalences. The default is those observed in the data. These should <br> add up to the actual number of people in the study at the start. |
| :--- | :--- |
| interp | Interpolation method for observed states, see prevalence.msm. <br> censtime <br> subset <br> covariates |
| Subject-specific maximum follow-up times, see prevalence.msm. <br> Vector of the subject identifiers to calculated observed prevalences for. <br> Covariate values for which to forecast expected state occupancy. See prevalence.msm <br> - if this function runs too slowly, as it may if there are continuous covariates, <br> replace covariates="population" with covariates="mean". |  |
| misccovariates | (Misclassification models only) Values of covariates on the misclassification <br> probability matrix. See prevalence.msm. |
| piecewise.times |  |$\quad$| Times at which piecewise-constant intensities change. See prevalence.msm. |
| :--- |

## Details

See prevalence.msm for details of the assumptions underlying this method.
Observed prevalences are plotted with a solid line, and expected prevalences with a dotted line.

## References

Gentleman, R.C., Lawless, J.F., Lindsey, J.C. and Yan, P. Multi-state Markov models for analysing incomplete disease history data with illustrations for HIV disease. Statistics in Medicine (1994) 13(3): 805-821.

## See Also

prevalence.msm

## Description

Plot a Kaplan-Meier estimate of the survival probability and compare it with the fitted survival probability from a msm model.

## Usage

```
## S3 method for class 'survfit.msm'
plot(x, from=1, to=NULL, range=NULL, covariates="mean",
    interp=c("start","midpoint"), ci=c("none","normal","bootstrap"), B=100,
                legend.pos=NULL, xlab="Time", ylab="Survival probability",
                lty=1, lwd=1, col="red", lty.ci=2, lwd.ci=1, col.ci="red",
                mark.time=TRUE, col.surv="blue", lty.surv=2, lwd.surv=1,
                survdata=FALSE,
                ...)
```


## Arguments

X
from
to Absorbing state to consider. Defaults to the highest-labelled absorbing state.
range Vector of two elements, giving the range of times to plot for.
covariates Covariate values for which to evaluate the expected probabilities. This can either be:
the string "mean", denoting the means of the covariates in the data (this is the default),
the number 0 , indicating that all the covariates should be set to zero,
or a list of values, with optional names. For example
list $(60,1)$
where the order of the list follows the order of the covariates originally given in the model formula, or a named list,
list (age $=60$, sex $=1$ )
but note the empirical curve is plotted for the full population. To consider subsets for the empirical curve, set survdata=TRUE to extract the survival data and build a survival plot by hand using plot. survfit.

| ci | If "none" (the default) no confidence intervals are plotted. If "normal" or "bootstrap", confidence intervals are plotted based on the respective method in pmatrix.msm. This is very computationally-intensive, since intervals must be computed at a series of times. |
| :---: | :---: |
| B | Number of bootstrap or normal replicates for the confidence interval. The default is 100 rather than the usual 1000, since these plots are for rough diagnostic purposes. |
| interp | If interp="start" (the default) then the entry time into the absorbing state is assumed to be the time it is first observed in the data. |
|  | If interp="midpoint" then the entry time into the absorbing state is assumed to be halfway between the time it is first observed and the previous observation time. This is generally more reasonable for "progressive" models with observations at arbitrary times. |
| legend.pos | Vector of the $x$ and $y$ position, respectively, of the legend. |
| xlab | x axis label. |
| ylab | y axis label. |
| lty | Line type for the fitted curve. See par. |
| lwd | Line width for the fitted curve. See par. |
| col | Colour for the fitted curve. See par. |
| lty.ci | Line type for the fitted curve confidence limits. See par. |
| lwd.ci | Line width for the fitted curve confidence limits. See par. |
| col.ci | Colour for the fitted curve confidence limits. See par. |
| mark.time | Mark the empirical survival curve at each censoring point, see lines.survfit. |
| col.surv | Colour for the empirical survival curve, passed to lines.survfit. See par. |
| lty.surv | Line type for the empirical survival curve, passed to lines.survfit. See par. |
| lwd.surv | Line width for the empirical survival curve, passed to lines. survfit. See par. |
| survdata | Set to TRUE to return the survival data frame constructed when plotting the empirical curve. This can be used for constructing survival plots by hand using plot.survfit. |
|  | Other arguments to be passed to the plot function which draws the fitted curve, or the lines.survfit function which draws the empirical curve. |

## Details

If the data represent observations of the process at arbitrary times, then the first occurrence of the absorbing state in the data will usually be greater than the actual first transition time to that state. Therefore the Kaplan-Meier estimate of the survival probability will be an overestimate.
The method of Turnbull (1976) could be used to give a non-parametric estimate of the time to an interval-censored event, and compared to the equivalent estimate from a multi-state model. This is implemented in the CRAN package interval (Fay and Shaw 2010).
This currently only handles time-homogeneous models.

## References

Turnbull, B. W. (1976) The empirical distribution function with arbitrarily grouped, censored and truncated data. J. R. Statist. Soc. B 38, 290-295.
Fay, MP and Shaw, PA (2010). Exact and Asymptotic Weighted Logrank Tests for Interval Censored Data: The interval R package. Journal of Statistical Software. http://www.jstatsoft.org/v36/ i02/. 36 (2):1-34.

## See Also

survfit, plot.survfit, plot.prevalence.msm
plotprog.msm Kaplan Meier estimates of incidence

## Description

Compute and plot Kaplan-Meier estimates of the probability that each successive state has not occurred yet.

## Usage

plotprog.msm(formula, subject, data, legend.pos=NULL, xlab="Time", ylab="1 - incidence probability", lwd=1, xlim=NULL, mark.time=TRUE, ...)

## Arguments

formula A formula giving the vectors containing the observed states and the corresponding observation times. For example,
state ~ time
Observed states should be in the set $1, \ldots, n$, where $n$ is the number of states.
subject Vector of subject identification numbers for the data specified by formula. If missing, then all observations are assumed to be on the same subject. These must be sorted so that all observations on the same subject are adjacent.
data An optional data frame in which the variables represented by state, time and subject can be found.
legend.pos Vector of the $x$ and $y$ position, respectively, of the legend.
$x l a b \quad x$ axis label.
ylab y axis label.
lwd Line width. See par.
$x \lim \quad x$ axis limits, e.g. $c(0,10)$ for an axis ranging from 0 to 10 . Default is the range of observation times.
mark.time Mark the empirical survival curve at each censoring point, see lines. survfit.
... Other arguments to be passed to the plot and lines. survfit functions.

## Details

If the data represent observations of the process at arbitrary times, then the first occurrence of the state in the data will usually be greater than the actual first transition time to that state. Therefore the probabilities plotted by plotprog.msm will be overestimates.

## See Also

```
survfit,plot.survfit
```

```
pmatrix.msm Transition probability matrix
```


## Description

Extract the estimated transition probability matrix from a fitted continuous-time multi-state model for a given time interval, at a given set of covariate values.

## Usage

pmatrix.msm(x=NULL, $t=1, \mathrm{t} 1=0$, covariates="mean", ci=c("none","normal","bootstrap"), cl=0.95, B=1000, cores=NULL, qmatrix=NULL, ...)

## Arguments

$x \quad$ A fitted multi-state model, as returned by msm.
$\mathrm{t} \quad$ The time interval to estimate the transition probabilities for, by default one unit.
t1 The starting time of the interval. Used for models $x$ with piecewise-constant intensities fitted using the pci option to msm. The probabilities will be computed on the interval $[\mathrm{t} 1, \mathrm{t} 1+\mathrm{t}]$.
covariates The covariate values at which to estimate the transition probabilities. This can either be:
the string "mean", denoting the means of the covariates in the data (this is the default),
the number 0 , indicating that all the covariates should be set to zero,
or a list of values, with optional names. For example
list $(60,1)$
where the order of the list follows the order of the covariates originally given in the model formula, or a named list,
list (age $=60$, sex $=1$ )
If some covariates are specified but not others, the missing ones default to zero.

For time-inhomogeneous models fitted using the pci option to msm, "covariates" here include only those specified using the covariates argument to msm, and exclude the artificial covariates representing the time period.
For time-inhomogeneous models fitted "by hand" by using a time-dependent covariate in the covariates argument to msm , the function pmatrix. piecewise.msm should be used to to calculate transition probabilities.
ci If "normal", then calculate a confidence interval for the transition probabilities by simulating $B$ random vectors from the asymptotic multivariate normal distribution implied by the maximum likelihood estimates (and covariance matrix) of the log transition intensities and covariate effects, then calculating the resulting transition probability matrix for each replicate. See, e.g. Mandel (2013) for a discussion of this approach.
If "bootstrap" then calculate a confidence interval by non-parametric bootstrap refitting. This is 1-2 orders of magnitude slower than the "normal" method, but is expected to be more accurate. See boot.msm for more details of bootstrapping in $\mathbf{m s m}$.
If "none" (the default) then no confidence interval is calculated.
cl Width of the symmetric confidence interval, relative to 1.
B
Number of bootstrap replicates, or number of normal simulations from the distribution of the MLEs
cores Number of cores to use for bootstrapping using parallel processing. See boot.msm for more details.
qmatrix A transition intensity matrix. Either this or a fitted model $\times$ must be supplied. No confidence intervals are available if qmatrix is supplied.
Optional arguments to be passed to MatrixExp to control the method of computing the matrix exponential.

## Details

For a continuous-time homogeneous Markov process with transition intensity matrix $Q$, the probability of occupying state $s$ at time $u+t$ conditionally on occupying state $r$ at time $u$ is given by the $(r, s)$ entry of the matrix $P(t)=\exp (t Q)$, where $\exp ()$ is the matrix exponential.
For non-homogeneous processes, where covariates and hence the transition intensity matrix $Q$ are piecewise-constant in time, the transition probability matrix is calculated as a product of matrices over a series of intervals, as explained in pmatrix.piecewise.msm.
The pmatrix.piecewise.msm function is only necessary for models fitted using a time-dependent covariate in the covariates argument to msm. For time-inhomogeneous models fitted using "pci", pmatrix.msm can be used, with arguments $t$ and $t$, to calculate transition probabilities over any time period.

## Value

The matrix of estimated transition probabilities $P(t)$ in the given time. Rows correspond to "fromstate" and columns to "to-state".

Or if ci="normal" or ci="bootstrap", pmatrix.msm returns a list with components estimates and ci, where estimates is the matrix of estimated transition probabilities, and ci is a list of two matrices containing the upper and lower confidence limits.

## Author(s)

C. H. Jackson <chris. jackson@mrc-bsu. cam.ac.uk>.

## References

Mandel, M. (2013). "Simulation based confidence intervals for functions with complicated derivatives." The American Statistician 67(2):76-81

## See Also

qmatrix.msm, pmatrix.piecewise.msm, boot.msm

pmatrix.piecewise.msm | Transition probability matrix for processes with piecewise-constant in- |
| :--- |
| tensities | tensities

## Description

Extract the estimated transition probability matrix from a fitted non-time-homogeneous multi-state model for a given time interval. This is a generalisation of pmatrix.msm to models with timedependent covariates. Note that pmatrix.msm is sufficient to calculate transition probabilities for time-inhomogeneous models fitted using the pci argument to msm.

## Usage

pmatrix.piecewise.msm(x=NULL, t1, t2, times, covariates, ci=c("none","normal","bootstrap"), cl=0.95, B=1000, cores=NULL, qlist=NULL, ...)

## Arguments

x
t1 The start of the time interval to estimate the transition probabilities for.
t2 The end of the time interval to estimate the transition probabilities for.
times Cut points at which the transition intensity matrix changes.
covariates A list with number of components one greater than the length of times. Each component of the list is specified in the same way as the covariates argument to pmatrix.msm. The components correspond to the covariate values in the intervals
(t1, times[1]],(times[1],times[2]],...,(times[length(times)],t2] (assuming that all elements of times are in the interval ( $\mathrm{t} 1, \mathrm{t} 2$ ) ).
$\left.\begin{array}{ll}\text { ci } & \begin{array}{l}\text { If "normal", then calculate a confidence interval for the transition probabilities } \\ \text { by simulating B random vectors from the asymptotic multivariate normal distri- } \\ \text { bution implied by the maximum likelihood estimates (and covariance matrix) of } \\ \text { the log transition intensities and covariate effects, then calculating the resulting } \\ \text { transition probability matrix for each replicate. }\end{array} \\ \text { If "bootstrap" then calculate a confidence interval by non-parametric bootstrap } \\ \text { refitting. This is 1-2 orders of magnitude slower than the "normal" method, but } \\ \text { is expected to be more accurate. See boot.msm for more details of bootstrapping } \\ \text { in msm. }\end{array}\right\}$

## Details

Suppose a multi-state model has been fitted, in which the transition intensity matrix $Q(x(t))$ is modelled in terms of time-dependent covariates $x(t)$. The transition probability matrix $P\left(t_{1}, t_{n}\right)$ for the time interval $\left(t_{1}, t_{n}\right)$ cannot be calculated from the estimated intensity matrix as $\exp \left(\left(t_{n}-t_{1}\right) Q\right)$, because $Q$ varies within the interval $t_{1}, t_{n}$. However, if the covariates are piecewise-constant, or can be approximated as piecewise-constant, then we can calculate $P\left(t_{1}, t_{n}\right)$ by multiplying together individual matrices $P\left(t_{i}, t_{i+1}\right)=\exp \left(\left(t_{i+1}-t_{i}\right) Q\right)$, calculated over intervals where Q is constant:

$$
P\left(t_{1}, t_{n}\right)=P\left(t_{1}, t_{2}\right) P\left(t_{2}, t_{3}\right) \ldots P\left(t_{n-1}, t_{n}\right)
$$

## Value

The matrix of estimated transition probabilities $P(t)$ for the time interval [t1, tn]. That is, the probabilities of occupying state $s$ at time $t_{n}$ conditionally on occupying state $r$ at time $t_{1}$. Rows correspond to "from-state" and columns to "to-state".

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

```
pmatrix.msm
```


## Examples

```
## Not run:
## In a clinical study, suppose patients are given a placebo in the
## first 5 weeks, then they begin treatment 1 at 5 weeks, and
## a combination of treatments 1 and 2 from 10 weeks.
## Suppose a multi-state model x has been fitted for the patients'
## progress, with treat1 and treat2 as time dependent covariates.
## Cut points for when treatment covariate changes
times <- c(0, 5, 10)
## Indicators for which treatments are active in the four intervals
## defined by the three cut points
covariates <- list( list (treat1=0, treat2=0), list (treat1=0, treat2=0), list(treat1=1, treat2=0),
list(treat1=1, treat2=1) )
## Calculate transition probabilities from the start of the study to 15 weeks
pmatrix.piecewise.msm(x, 0, 15, times, covariates)
## End(Not run)
```

    pnext.msm Probability of each state being next
    
## Description

Compute a matrix of the probability of each state $s$ being the next state of the process after each state $r$. Together with the mean sojourn times in each state (sojourn.msm), these fully define a continuous-time Markov model.

## Usage

pnext.msm(x, covariates = "mean",
ci=c("normal","bootstrap","delta","none"), cl = 0.95,
$B=1000$, cores=NULL)

## Arguments

$x \quad$ A fitted multi-state model, as returned by msm.
covariates The covariate values at which to estimate the intensities. This can either be:
the string "mean", denoting the means of the covariates in the data (this is the default),
the number 0 , indicating that all the covariates should be set to zero,
or a list of values, with optional names. For example
list $(60,1)$
where the order of the list follows the order of the covariates originally given in the model formula, or a named list,
list (age $=60$, sex $=1$ )
ci If "normal" (the default) then calculate a confidence interval by simulating B random vectors from the asymptotic multivariate normal distribution implied by the maximum likelihood estimates (and covariance matrix) of the log transition intensities and covariate effects, then transforming.
If "bootstrap" then calculate a confidence interval by non-parametric bootstrap refitting. This is 1-2 orders of magnitude slower than the "normal" method, but is expected to be more accurate. See boot.msm for more details of bootstrapping in $\mathbf{m s m}$.
If "delta" then confidence intervals are calculated based on the delta method SEs of the log rates, but this is not recommended since it may not respect the constraint that probabilities are less than one.
cl Width of the symmetric confidence interval to present. Defaults to 0.95 .
B Number of bootstrap replicates, or number of normal simulations from the distribution of the MLEs.
cores Number of cores to use for bootstrapping using parallel processing. See boot.msm for more details.

## Details

For a continuous-time Markov process in state $r$, the probability that the next state is $s$ is $-q_{r s} / q_{r r}$, where $q_{r s}$ is the transition intensity (qmatrix.msm).

A continuous-time Markov model is fully specified by these probabilities together with the mean sojourn times $-1 / q_{r r}$ in each state $r$. This gives a more intuitively meaningful description of a model than the intensity matrix.
Remember that msm deals with continuous-time, not discrete-time models, so these are not the same as the probability of observing state $s$ at a fixed time in the future. Those probabilities are given by pmatrix.msm.

## Value

The matrix of probabilities that the next move of a process in state $r$ (rows) is to state $s$ (columns).

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

qmatrix.msm,pmatrix.msm,qratio.msm

## Description

Probabilities of having visited each state by a particular time in a continuous time Markov model.

## Usage

ppass.msm(x=NULL, qmatrix=NULL, tot, start="all", covariates="mean", piecewise.times=NULL, piecewise.covariates=NULL, ci=c("none","normal","bootstrap"), cl=0.95, B=1000, cores=NULL, ...)

## Arguments

$x \quad$ A fitted multi-state model, as returned by msm.
qmatrix Instead of $x$, you can simply supply a transition intensity matrix in qmatrix.
tot Finite time to forecast the passage probabilites for.
start Starting state (integer). By default (start="all"), this will return a matrix one row for each starting state.
Alternatively, this can be used to obtain passage probabilities from a set of states, rather than single states. To achieve this, state is set to a vector of weights, with length equal to the number of states in the model. These weights should be proportional to the probability of starting in each of the states in the desired set, so that weights of zero are supplied for other states. The function will calculate the weighted average of the passage probabilities from each of the corresponding states.
covariates Covariate values defining the intensity matrix for the fitted model x , as supplied to qmatrix.msm.
piecewise.times
Currently ignored: not implemented for time-inhomogeneous models.
piecewise.covariates
Currently ignored: not implemented for time-inhomogeneous models.
ci If "normal", then calculate a confidence interval by simulating B random vectors from the asymptotic multivariate normal distribution implied by the maximum likelihood estimates (and covariance matrix) of the log transition intensities and covariate effects.
If "bootstrap" then calculate a confidence interval by non-parametric bootstrap refitting. This is 1-2 orders of magnitude slower than the "normal" method, but is expected to be more accurate. See boot.msm for more details of bootstrapping in $\mathbf{~ m s m}$.
If "none" (the default) then no confidence interval is calculated.
cl Width of the symmetric confidence interval, relative to 1.

B Number of bootstrap replicates.
cores Number of cores to use for bootstrapping using parallel processing. See boot.msm for more details.
... Arguments to pass to MatrixExp.

## Details

The passage probabilities to state $s$ are computed by setting the $s$ th row of the transition intensity matrix $Q$ to zero, giving an intensity matrix $Q *$ for a simplified model structure where state $s$ is absorbing. The probabilities of passage are then equivalent to row $s$ of the transition probability matrix $\operatorname{Exp}(t Q *)$ under this simplified model for $t=$ tot.
Note this is different from the probability of occupying each state at exactly time $t$, given by pmatrix.msm. The passage probability allows for the possibility of having visited the state before $t$, but then occupying a different state at $t$.
The mean of the passage distribution is the expected first passage time, efpt.msm.
This function currently only handles time-homogeneous Markov models. For time-inhomogeneous models the covariates are held constant at the value supplied, by default the column means of the design matrix over all observations.

## Value

A matrix whose $r, s$ entry is the probability of having visited state $s$ at least once before time $t$, given the state at time 0 is $r$. The diagonal entries should all be 1 .

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## References

Norris, J. R. (1997) Markov Chains. Cambridge University Press.

## See Also

efpt.msm, totlos.msm, boot.msm.

## Examples

```
Q <- rbind(c(-0.5, 0.25, 0, 0.25), c(0.166, -0.498, 0.166, 0.166),
    c(0, 0.25, -0.5, 0.25), c(0, 0, 0, 0))
## ppass[1,2](t) converges to 0.5 with t, since given in state 1, the
## probability of going to the absorbing state 4 before visiting state
## 2 is 0.5, and the chance of still being in state 1 at t decreases.
ppass.msm(qmatrix=Q, tot=2)
ppass.msm(qmatrix=Q, tot=20)
ppass.msm(qmatrix=Q, tot=100)
```

```
Q <- Q[1:3,1:3]; diag(Q) <- 0; diag(Q) <- -rowSums(Q)
## Probability of about 1/2 of visiting state 3 by time 10.5, the
## median first passage time
ppass.msm(qmatrix=Q, tot=10.5)
## Mean first passage time from state 2 to state 3 is 10.02: similar
## to the median
efpt.msm(qmatrix=Q, tostate=3)
```

prevalence.msm Tables of observed and expected prevalences

## Description

This provides a rough indication of the goodness of fit of a multi-state model, by estimating the observed numbers of individuals occupying each state at a series of times, and comparing these with forecasts from the fitted model.

## Usage

prevalence.msm(x, times=NULL, timezero=NULL, initstates=NULL, covariates="population", misccovariates="mean", piecewise.times=NULL, piecewise.covariates=NULL, ci=c("none","normal","bootstrap"), cl=0.95, B=1000, cores=NULL, interp=c("start","midpoint"), censtime=Inf, subset=NULL, plot=FALSE, ...)

## Arguments

X
times Series of times at which to compute the observed and expected prevalences of states.
timezero Initial time of the Markov process. Expected values are forecasted from here. Defaults to the minimum of the observation times given in the data.
initstates Optional vector of the same length as the number of states. Gives the numbers of individuals occupying each state at the initial time, to be used for forecasting expected prevalences. The default is those observed in the data. These should add up to the actual number of people in the study at the start.
covariates Covariate values for which to forecast expected state occupancy. With the default covariates="population", expected prevalences are produced by summing model predictions over the covariates observed in the original data, for a fair comparison with the observed prevalences. This may be slow, particularly with continuous covariates.
Predictions for fixed covariates can be obtained by supplying covariate values in the standard way, as in qmatrix.msm. Therefore if covariates="population"
is too slow, using the mean observed values through covariates="mean" may give a reasonable approximation.
This argument is ignored if piecewise. times is specified. If there are a mixture of time-constant and time-dependent covariates, then the values for all covariates should be supplied in piecewise.covariates.
misccovariates (Misclassification models only) Values of covariates on the misclassification probability matrix for converting expected true to expected misclassified states. Ignored if covariates="population", otherwise defaults to the mean values of the covariates in the data set.

## piecewise.times

Times at which piecewise-constant intensities change. See pmatrix. piecewise.msm for how to specify this. Ignored if covariates="population". This is only required for time-inhomogeneous models specified using explicit time-dependent covariates, and should not be used for models specified using "pci".
piecewise.covariates
Covariates on which the piecewise-constant intensities depend. See pmatrix. piecewise.msm for how to specify this. Ignored if covariates="population".
ci If "normal", then calculate a confidence interval for the expected prevalences by simulating $B$ random vectors from the asymptotic multivariate normal distribution implied by the maximum likelihood estimates (and covariance matrix) of the log transition intensities and covariate effects, then calculating the expected prevalences for each replicate.
If "bootstrap" then calculate a confidence interval by non-parametric bootstrap refitting. This is 1-2 orders of magnitude slower than the "normal" method, but is expected to be more accurate. See boot.msm for more details of bootstrapping in $\mathbf{m s m}$.
If "none" (the default) then no confidence interval is calculated.
cl Width of the symmetric confidence interval, relative to 1
B Number of bootstrap replicates
cores Number of cores to use for bootstrapping using parallel processing. See boot.msm for more details.
interp $\quad$ Suppose an individual was observed in states $S_{r-1}$ and $S_{r}$ at two consecutive times $t_{r-1}$ and $t_{r}$, and we want to estimate 'observed' prevalences at a time $t$ between $t_{r-1}$ and $t_{r}$.
If interp="start", then individuals are assumed to be in state $S_{r-1}$ at time $t$, the same state as they were at $t_{r-1}$.
If interp="midpoint" then if $t<=\left(t_{r-1}+t_{r}\right) / 2$, the midpoint of $t_{r-1}$ and $t_{r}$, the state at $t$ is assumed to be $S_{r-1}$, otherwise $S_{r}$. This is generally more reasonable for "progressive" models.
censtime Adjustment to the observed prevalences to account for limited follow-up in the data.
If the time is greater than censtime and the patient has reached an absorbing state, then that subject will be removed from the risk set. For example, if patients have died but would only have been observed up to this time, then this avoids overestimating the proportion of people who are dead at later times.

This can be supplied as a single value, or as a vector with one element per subject (after any subset has been taken), in the same order as the original data. This vector also only includes subjects with complete data, thus it excludes for example subjects with only one observation (thus no observed transitions), and subjects for whom every observation has missing values. (Note, to help construct this, the complete data used for the model fit can be accessed with model.frame ( x ), where x is the fitted model object)
This is ignored if it is less than the subject's maximum observation time.
subset Subset of subjects to calculate observed prevalences for.
plot Generate a plot of observed against expected prevalences. See plot.prevalence.msm
... Further arguments to pass to plot.prevalence.msm.

## Details

The fitted transition probability matrix is used to forecast expected prevalences from the state occupancy at the initial time. To produce the expected number in state $j$ at time $t$ after the start, the number of individuals under observation at time $t$ (including those who have died, but not those lost to follow-up) is multiplied by the product of the proportion of individuals in each state at the initial time and the transition probability matrix in the time interval $t$. The proportion of individuals in each state at the "initial" time is estimated, if necessary, in the same way as the observed prevalences.
For misclassification models (fitted using an ematrix), this aims to assess the fit of the full model for the observed states. That is, the combined Markov progression model for the true states and the misclassification model. Thus, expected prevalences of true states are estimated from the assumed proportion occupying each state at the initial time using the fitted transition probabiliy matrix. The vector of expected prevalences of true states is then multiplied by the fitted misclassification probability matrix to obtain the expected prevalences of observed states.
For general hidden Markov models, the observed state is taken to be the predicted underlying state from the Viterbi algorithm (viterbi.msm). The goodness of fit of these states to the underlying Markov model is tested.
In any model, if there are censored states, then these are replaced by imputed values of highest probability from the Viterbi algorithm in order to calculate the observed state prevalences.
For an example of this approach, see Gentleman et al. (1994).

## Value

A list of matrices, with components:
Observed Table of observed numbers of individuals in each state at each time
Observed percentages
Corresponding percentage of the individuals at risk at each time.
Expected Table of corresponding expected numbers.
Expected percentages
Corresponding percentage of the individuals at risk at each time.
Or if ci.boot = TRUE, the component Expected is a list with components estimates and ci. estimates is a matrix of the expected prevalences, and ci is a list of two matrices, containing the confidence limits. The component Expected percentages has a similar format.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## References

Gentleman, R.C., Lawless, J.F., Lindsey, J.C. and Yan, P. Multi-state Markov models for analysing incomplete disease history data with illustrations for HIV disease. Statistics in Medicine (1994) 13(3): 805-821.
Titman, A.C., Sharples, L. D. Model diagnostics for multi-state models. Statistical Methods in Medical Research (2010) 19(6):621-651.

## See Also

msm, summary.msm
print.msm Print a fitted msm model object

## Description

Print a fitted msm model object

## Usage

\#\# S3 method for class 'msm'
print(x, covariates=NULL, digits=4, ...)
printnew.msm(x, covariates=NULL, digits=4, ...)

## Arguments

x
Output from msm, representing a fitted multi-state model object.
covariates Covariates for which to print "baseline" transition intensities or misclassification probabilities. See qmatrix.msm for more details.
digits Minimum number of significant digits, passed to format. Defaults to 4.
... Other arguments to be passed to format.

## Details

This is the new method of formatting msm objects for printing. The old method was based on printing lists of matrices. That produced a lot of wasted space for parameters which were zero, and it was difficult to match corresponding numbers between matrices. The new method presents all the transition intensities and covariate effects as a single compact table, and likewise for misclassification matrices.
Also in the old method, covariate effects were presented as log hazard ratios or log odds ratios. The $\log$ scale is more convenient mathematically, but unnatural to interpret. The new method presents hazard ratios for covariates on transition intensities and odds ratios for misclassification probabilities.
printnew.msm is an alias for print.msm.

## Value

The object returned by print.msm is a numeric matrix with one column for each estimate or confidence limit for intensities and their covariates, in the same arrangement as printed, but with the underlying numbers in full precision. The results formatted for printing are stored in the "formatted" attribute of the object, as a character matrix. These can alternatively be produced by msm.form.qoutput, which has no printing side-effect. msm.form. eoutput produces the same arrangement for misclassification probabilities instead of intensities.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

msm, printold.msm, msm.form.qoutput.

```
printold.msm Print a fitted msm model object
```


## Description

Print a fitted msm model object (in old format, from msm 1.3.1 and earlier)

## Usage

printold.msm(x,...)

## Arguments

x
Output from msm, representing a fitted multi-state model object.
... Other arguments to be passed to format.

## Details

See print.msm for a better and cleaner output format, and an explanation of the change.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

```
print.msm
```

psor Psoriatic arthritis data

## Description

A series of observations of grades of psoriatic arthritis, as indicated by numbers of damaged joints.

## Usage

psor

## Format

A data frame containing 806 observations, representing visits to a psoriatic arthritis (PsA) clinic from 305 patients. The rows are grouped by patient number and ordered by examination time. Each row represents an examination and contains additional covariates.
\(\left.$$
\begin{array}{rll}\text { ptnum } & \text { (numeric) } & \begin{array}{l}\text { Patient identification number } \\
\text { months }\end{array}
$$ <br>

(numeric) \& Examination time in months\end{array}\right\}\)| state | (numeric) |
| :--- | :--- |
|  |  |
| Clinical state of PsA. Patients in states 1,2,3 and 4 |  |
| have 0,1 to 4, 5 to 9 and 10 or more damaged joints, |  |
| respectively. |  |

## References

Gladman, D. D. and Farewell, V.T. (1999) Progression in psoriatic arthritis: role of time-varying clinical indicators. J. Rheumatol. 26(11):2409-13

## Examples

```
## Four-state progression-only model with high effusion and low
## sedimentation rate as covariates on the progression rates. High
## effusion is assumed to have the same effect on the 1-2, 2-3, and 3-4
## progression rates, while low sedimentation rate has the same effect
## on the 1-2 and 2-3 intensities, but a different effect on the 3-4.
data(psor)
psor.q <- rbind(c(0,0.1,0,0),c(0,0,0.1,0),c(0,0,0,0.1),c(0,0,0,0))
psor.msm <- msm(state ~ months, subject=ptnum, data=psor,
    qmatrix = psor.q, covariates = ~ollwsdrt+hieffusn,
    constraint = list(hieffusn=c(1,1,1),ollwsdrt=c(1,1,2)),
    fixedpars=FALSE, control = list(REPORT=1,trace=2), method="BFGS")
qmatrix.msm(psor.msm)
sojourn.msm(psor.msm)
hazard.msm(psor.msm)
```


## Description

A list representing the model for covariates on transition intensities

## Value

| npars | Number of covariate effect parameters. This is defined as the number of covari- <br> ates on intensities (with factors expanded as contrasts) multiplied by the number <br> of allowed transitions in the model. <br> Note if msm was called with covariates set to a list of different covariates for <br> different intensities, then this will include covariate effects that are implicitly <br> defined as zero by this list. The information in paramdata objects can be used <br> to identify wich ones are fixed at zero. <br> This also includes any timeperiod covariates in a time-inhomogeneous model <br> defined by the pci option to msm. |
| :--- | :--- |
| ndpars | Number of distinct covariate effect parameters, as npars, but after any equality <br> constraints have been applied. |
| ncovs | Number of covariates on intensities, with factors expanded as contrasts. |
| constr | List of equality constraints on these covariate effects, as supplied in the constraint <br> argument to msm. |
| covlabels | Names / labels of these covariates in the model matrix (see model.matrix.msm). |
| covmeans | Initial values for these covariate effects, as a vector formed from the covinits <br> list supplied to msm. |
| Means of these covariates in the data (excluding data not required to fit the <br> model, such as observations with missing data in other elements or subjects' last <br> observations). This includes means of 0/1 factor contrasts as well as continuous |  |
| covariates (for historic reasons, which may not be sensible). |  |

## See Also

msm.object.
qgeneric Generic function to find quantiles of a distribution

## Description

Generic function to find the quantiles of a distribution, given the equivalent probability distribution function.

## Usage

qgeneric(pdist, p, special=NULL, ...)

## Arguments

pdist Probability distribution function, for example, pnorm for the normal distribution, which must be defined in the current workspace. This should accept and return vectorised parameters and values. It should also return the correct values for the entire real line, for example a positive distribution should have pdist $(x)==0$ for $x<0$.
p Vector of probabilities to find the quantiles for.
special Vector of character strings naming arguments of the distribution function that should not be vectorised over. Used, for example, for the rate and $t$ arguments in qpexp.
... The remaining arguments define parameters of the distribution pdist. These MUST be named explicitly.
This may also contain the standard arguments log.p (logical; default FALSE, if TRUE, probabilities $p$ are given as $\log (p)$ ), and lower. tail (logical; if TRUE (default), probabilities are $\mathrm{P}[\mathrm{X}<=\mathrm{x}]$ otherwise, $\mathrm{P}[\mathrm{X}>\mathrm{x}]$.).
If the distribution is bounded above or below, then this should contain arguments lbound and ubound respectively, and these will be returned if $p$ is 0 or 1 respectively. Defaults to -Inf and Inf respectively.

## Details

This function is intended to enable users to define " $q$ " functions for new distributions, in cases where the distribution function pdist is available analytically, but the quantile function is not.
It works by finding the root of the equation $h(q)=p \operatorname{dist}(q)-p=0$. Starting from the interval $(-1,1)$, the interval width is expanded by $50 \%$ until $h()$ is of opposite sign at either end. The root is then found using uniroot.

This assumes a suitably smooth, continuous distribution.
An identical function is provided in the flexsurv package.

## Value

Vector of quantiles of the distribution at p .

## Author(s)

Christopher Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## Examples

qnorm $(c(0.025,0.975), 0,1)$
qgeneric(pnorm, $c(0.025,0.975)$, mean $=0, \mathrm{sd}=1)$ \# must name the arguments

Transition intensity matrix

## Description

Extract the estimated transition intensity matrix, and the corresponding standard errors, from a fitted multi-state model at a given set of covariate values.

## Usage

qmatrix.msm(x, covariates="mean", sojourn=FALSE,
ci=c("delta","normal","bootstrap","none"), cl=0.95, $\mathrm{B}=1000$, cores=NULL)

## Arguments

$x \quad$ A fitted multi-state model, as returned by msm.
covariates The covariate values at which to estimate the intensity matrix. This can either be:
the string "mean", denoting the means of the covariates in the data (this is the default),
the number 0 , indicating that all the covariates should be set to zero,
or a list of values, with optional names. For example
list $(60,1)$
where the order of the list follows the order of the covariates originally given in the model formula. Or more clearly, a named list,
list (age $=60$, sex $=1$ )
If some covariates are specified but not others, the missing ones default to zero.
With covariates="mean", for factor / categorical variables, the mean of the $0 / 1$ dummy variable for each factor level is used, representing an average over all values in the data, rather than a specific factor level.
sojourn Set to TRUE if the estimated sojourn times and their standard errors should also be returned.

$$
\begin{array}{ll}
\text { ci } & \text { If "delta" (the default) then confidence intervals are calculated by the delta } \\
\text { method, or by simple transformation of the Hessian in the very simplest cases. } \\
\text { Normality on the log scale is assumed. } \\
\text { If "normal", then calculate a confidence interval by simulating B random vectors } \\
\text { from the asymptotic multivariate normal distribution implied by the maximum } \\
\text { likelihood estimates (and covariance matrix) of the log transition intensities and } \\
\text { covariate effects, then transforming. } \\
\text { If "bootstrap" then calculate a confidence interval by non-parametric bootstrap } \\
\text { refitting. This is 1-2 orders of magnitude slower than the "normal" method, but } \\
\text { is expected to be more accurate. See boot.msm for more details of bootstrapping } \\
\text { in msm. }
\end{array} \quad \begin{aligned}
& \text { Width of the symmetric confidence interval to present. Defaults to } 0.95 . \\
& \text { cl } \\
& \text { Bumber of bootstrap replicates, or number of normal simulations from the dis- } \\
& \text { cores }
\end{aligned} \quad \begin{aligned}
& \text { tribution of the MLEs. }
\end{aligned}
$$

## Details

Transition intensities and covariate effects are estimated on the $\log$ scale by msm. A covariance matrix is estimated from the Hessian of the maximised log-likelihood.

A more practically meaningful parameterisation of a continuous-time Markov model with transition intensities $q_{r s}$ is in terms of the mean sojourn times $-1 / q_{r r}$ in each state $r$ and the probabilities that the next move of the process when in state $r$ is to state $s,-q_{r s} / q_{r r}$.

## Value

A list with components:

| estimate | Estimated transition intensity matrix. |
| :--- | :--- |
| SE | Corresponding approximate standard errors. |
| L | Lower confidence limits |
| U | Upper confidence limits |

Or if ci="none", then qmatrix.msm just returns the estimated transition intensity matrix.
If sojourn is TRUE, extra components called sojourn, sojournSE, sojournL and sojournU are included, containing the estimates, standard errors and confidence limits, respectively, of the mean sojourn times in each transient state.
The default print method for objects returned by qmatrix.msm presents estimates and confidence limits. To present estimates and standard errors, do something like
qmatrix.msm(x)[c("estimates","SE")]

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

pmatrix.msm, sojourn.msm, deltamethod, ematrix.msm
qmodel. object Developer documentation: transition model structure object

## Description

A list giving information about the structure of states and allowed transitions in a multi-state model, and options for likelihood calculation. Used in internal computations, and returned in a fitted msm model object.

## Value

| nstates | Number of states |
| :--- | :--- |
| iso | Label for which basic structure the model is isomorphic to in the list of structures <br> for which analytic formulae for the transition probabilities are implemented <br> in the source file src/analyticp.c. This list is given by the internal ob- <br> ject msm: : : msm.graphs which is defined and documented in the source file <br> R/constants.R. <br> iso is 0 if the analytic P matrix is not implemented for this structure, or if <br> analytic P matrix calculations are disabled using use. analyticp=FALSE in the <br> call to msm. |
| perm | Permutation required to convert the base isomorphism into the structure of this <br> model. A vector of integers whose $r$ th element is the state number in the base <br> structure representing state $r$ in the current structure. |
| qperm | Inverse permutation: vector whose $r$ th element is the state number in the current <br> structure representing the $r$ th state in the base structure. |
| npars | Number of allowed instantaneous transitions, equal to sum(imatrix). <br> Indicator matrix for allowed instantaneous transitions. This has $(r, s)$ entry 1 if <br> the transition from $r$ to $s$ is permitted in continuous time, and 0 otherwise. The <br> diagonal entries are arbitrarily set to 0. |
| qmatrix | Matrix of initial values for the transition intensities, supplied as the qmatrix <br> argument of msm. |
| inits | Vector of these initial values, reading across rows of qmatrix and excluding the <br> diagonal and disallowed transitions. |
| constr | Indicators for equality constraints on baseline intensities, taken from the qconstraint <br> argument to msm, and mapped if necessary to the set (1,2,3,...). |
| ndpars | Number of distinct allowed instantaneous transitions, after applying equality <br> constraints. <br> Use expm package to calculate matrix exponentials for likelihoods, as supplied <br> to the use.expm argument of msm. TRUE or FALSE. |
| expm |  |

## See Also

msm. object,emodel.object, hmodel.object.

## Description

Compute the estimate and approximate standard error of the ratio of two estimated transition intensities from a fitted multi-state model at a given set of covariate values.

## Usage

qratio.msm(x, ind1, ind2, covariates = "mean",
ci=c("delta","normal","bootstrap","none"), cl = 0.95,
B=1000, cores=NULL)

## Arguments

X
ind1 Pair of numbers giving the indices in the intensity matrix of the numerator of the ratio, for example, $c(1,2)$.
ind2 Pair of numbers giving the indices in the intensity matrix of the denominator of the ratio, for example, $c(2,1)$.
covariates The covariate values at which to estimate the intensities. This can either be:
the string "mean", denoting the means of the covariates in the data (this is the default),
the number 0 , indicating that all the covariates should be set to zero,
or a list of values, with optional names. For example
list $(60,1)$
where the order of the list follows the order of the covariates originally given in the model formula, or a named list,
list (age $=60$, sex $=1$ )
ci If "delta" (the default) then confidence intervals are calculated by the delta method.
If "normal", then calculate a confidence interval by simulating $B$ random vectors from the asymptotic multivariate normal distribution implied by the maximum likelihood estimates (and covariance matrix) of the log transition intensities and covariate effects, then transforming.
If "bootstrap" then calculate a confidence interval by non-parametric bootstrap refitting. This is 1-2 orders of magnitude slower than the "normal" method, but is expected to be more accurate. See boot.msm for more details of bootstrapping in $\mathbf{~ m s m}$.
cl Width of the symmetric confidence interval to present. Defaults to 0.95 .
$\begin{array}{ll}\text { B } & \begin{array}{l}\text { Number of bootstrap replicates, or number of normal simulations from the dis- } \\ \text { tribution of the MLEs }\end{array} \\ \text { cores } & \begin{array}{l}\text { Number of cores to use for bootstrapping using parallel processing. See boot.msm } \\ \text { for more details. }\end{array}\end{array}$

## Details

For example, we might want to compute the ratio of the progression rate and recovery rate for a fitted model disease.msm with a health state (state 1) and a disease state (state 2). In this case, the progression rate is the $(1,2)$ entry of the intensity matrix, and the recovery rate is the $(2,1)$ entry. Thus to compute this ratio with covariates set to their means, we call qratio.msm(disease.msm, $c(1,2), c(2,1))$.
Standard errors are estimated by the delta method. Confidence limits are estimated by assuming normality on the log scale.

## Value

A named vector with elements estimate, se, $L$ and $U$ containing the estimate, standard error, lower and upper confidence limits, respectively, of the ratio of intensities.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

qmatrix.msm
recreate.olddata Convert data stored in msm object to old format

## Description

Converts the data element of msm objects to the old format.

## Usage

recreate.olddata(x)

## Arguments

x
Object returned by the msm function, representing a fitted multi-state model.

## Details

This is just provided for convenience and to illustrate the changes. It is not guaranteed to be complete, and is liable to be withdrawn. Users who were relying on the previous undocumented format are advised to upgrade their code to use the new format, which uses model frames and model design matrices in the standard format used in version 1.4, based on model. frame and model.matrix.

## Value

A list of vectors and matrices in the undocumented ad-hoc format used for the data component of msm objects in $\mathbf{~ m s m}$ versions 1.3.1 and earlier.

```
scoreresid.msm Score residuals
```


## Description

Score residuals for detecting outlying subjects.

## Usage

scoreresid.msm(x, plot=FALSE)

## Arguments

$x \quad$ A fitted multi-state model, as returned by msm.
plot If TRUE, display a simple plot of the residuals in subject order, labelled by subject identifiers

## Details

The score residual for a single subject is

$$
U(\theta)^{T} I(\theta)^{-1} U(\theta)
$$

where $U(\theta)$ is the vector of first derivatives of the log-likelihood for that subject at maximum likelihood estimates $\theta$, and $I(\theta)$ is the observed Fisher information matrix, that is, the matrix of second derivatives of minus the log-likelihood for that subject at theta.
Subjects with a higher influence on the maximum likelihood estimates will have higher score residuals.

These are only available for models with analytic derivatives (which includes all non-hidden and most hidden Markov models).

## Value

Vector of the residuals, named by subject identifiers.

## Author(s)

Andrew Titman [a.titman@lancaster.ac.uk](mailto:a.titman@lancaster.ac.uk) (theory), Chris Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk) (code)

Simulate one individual trajectory from a continuous-time Markov model

## Description

Simulate one realisation from a continuous-time Markov process up to a given time.

## Usage

```
sim.msm(qmatrix, maxtime, covs=NULL, beta=NULL, obstimes=0, start=1,
mintime=0)
```


## Arguments

qmatrix The transition intensity matrix of the Markov process. The diagonal of qmatrix is ignored, and computed as appropriate so that the rows sum to zero. For example, a possible qmatrix for a three state illness-death model with recovery is:
rbind( $c(0,0.1,0.02), c(0.1,0,0.01), c(0,0,0))$
maxtime Maximum time for the simulated process.
covs Matrix of time-dependent covariates, with one row for each observation time and one column for each covariate.
beta Matrix of linear covariate effects on log transition intensities. The rows correspond to different covariates, and the columns to the transition intensities. The intensities are ordered by reading across rows of the intensity matrix, starting with the first, counting the positive off-diagonal elements of the matrix.
obstimes Vector of times at which the covariates are observed.
start $\quad$ Starting state of the process. Defaults to 1.
mintime $\quad$ Starting time of the process. Defaults to 0.

## Details

The effect of time-dependent covariates on the transition intensity matrix for an individual is determined by assuming that the covariate is a step function which remains constant in between the individual's observation times.

## Value

A list with components,

| states | Simulated states through which the process moves. This ends with either an <br> absorption before obstime, or a transient state at obstime. |
| :--- | :--- |
| times | Exact times at which the process changes to the corresponding states |
| qmatrix | The given transition intensity matrix |

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

simmulti.msm

## Examples

```
qmatrix <- rbind(
    c(-0.2, 0.1, 0.1 ),
    c(0.5, -0.6, 0.1),
    c(0, 0, 0)
    )
sim.msm(qmatrix, 30)
```

simfitted.msm
Simulate from a Markov model fitted using msm

## Description

Simulate a dataset from a Markov model fitted using msm, using the maximum likelihood estimates as parameters, and the same observation times as in the original data.

## Usage

simfitted.msm(x, drop.absorb=TRUE, drop.pci.imp=TRUE)

## Arguments

x
drop.absorb Should repeated observations in an absorbing state be omitted. Use the default of TRUE to avoid warnings when using the simulated dataset for further msm fits. Or set to FALSE if exactly the same number of observations as the original data are needed.
drop.pci.imp In time-inhomogeneous models fitted using the pci option to msm, censored observations are inserted into the data by msm at the times where the intensity changes, but dropped by default when simulating from the fitted model using this function. Set this argument to FALSE to keep these observations and the corresponding indicator variable.

## Details

This function is a wrapper around simmulti.msm, and only simulates panel-observed data. To generate datasets with the exact times of transition, use the lower-level sim.msm.
Markov models with misclassified states fitted through the ematrix option to msm are supported, but not general hidden Markov models with hmodel. For misclassification models, this function includes misclassification in the simulated states.
This function is used for parametric bootstrapping to estimate the null distribution of the test statistic in pearson.msm.

## Value

A dataset with variables as described in simmulti.msm.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

```
simmulti.msm, sim.msm, pearson.msm, msm.
```

simmulti.msm Simulate multiple trajectories from a multi-state Markov model with arbitrary observation times

## Description

Simulate a number of individual realisations from a continuous-time Markov process. Observations of the process are made at specified arbitrary times for each individual, giving panel-observed data.

## Usage

```
simmulti.msm(data, qmatrix, covariates=NULL, death = FALSE, start,
                    ematrix=NULL, misccovariates=NULL, hmodel=NULL, hcovariates=NULL,
                        censor.states=NULL, drop.absorb=TRUE)
```


## Arguments

data A data frame with a mandatory column named time, representing observation times. The optional column named subject, corresponds to subject identification numbers. If not given, all observations are assumed to be on the same individual. Observation times should be sorted within individuals. The optional column named cens indicates the times at which simulated states should be censored. If cens $==0$ then the state is not censored, and if cens==k, say, then all simulated states at that time which are in the set censor. states are replaced by k. Other named columns of the data frame represent any covariates, which may be time-constant or time-dependent. Time-dependent covariates are assumed to be constant between the observation times.

| qmatrix | The transition intensity matrix of the Markov process, with any covariates set to zero. The diagonal of qmatrix is ignored, and computed as appropriate so that the rows sum to zero. For example, a possible qmatrix for a three state illness-death model with recovery is: <br> rbind ( $c(0,0.1,0.02), c(0.1,0,0.01), c(0,0,0))$ |
| :---: | :---: |
| covariates | List of linear covariate effects on log transition intensities. Each element is a vector of the effects of one covariate on all the transition intensities. The intensities are ordered by reading across rows of the intensity matrix, starting with the first, counting the positive off-diagonal elements of the matrix. <br> For example, for a multi-state model with three transition intensities, and two covariates $x$ and $y$ on each intensity, <br> covariates $=1$ ist $(x=c(-0.3,-0.3,-0.3), y=c(0.1,0.1,0.1))$ |
| death | Vector of indices of the death states. A death state is an absorbing state whose time of entry is known exactly, but the individual is assumed to be in an unknown transient state ("alive") at the previous instant. This is the usual situation for times of death in chronic disease monitoring data. For example, if you specify death $=c(4,5)$ then states 4 and 5 are assumed to be death states. <br> death $=$ TRUE indicates that the final state is a death state, and death $=$ FALSE (the default) indicates that there is no death state. |
| start | A vector with the same number of elements as there are distinct subjects in the data, giving the states in which each corresponding individual begins. Or a single number, if all of these are the same. Defaults to state 1 for each subject. |
| ematrix | An optional misclassification matrix for generating observed states conditionally on the simulated true states. As defined in msm. |
| misccovariates | Covariate effects on misclassification probabilities via multinomial logistic regression. Linear effects operate on the $\log$ of each probability relative to the probability of classification in the correct state. In same format as covariates. |
| hmodel | An optional hidden Markov model for generating observed outcomes conditionally on the simulated true states. As defined in msm. |
| hcovariates | List of the same length as hmodel, defining any covariates governing the hidden Markov outcome models. Unlike in the msm function, this should also define the values of the covariate effects. Each element of the list is a named vector of the initial values for each set of covariates for that state. For example, for a three-state hidden Markov model with two, one and no covariates on the state 1, 2 and 3 outcome models respectively, <br> hcovariates $=$ list $(c$ (acute $=-8$, age $=0), c($ acute $=-8)$, NULL) |

censor.states Set of simulated states which should be replaced by a censoring indicator at censoring times. By default this is all transient states (representing alive, with unknown state).
drop. absorb Drop repeated observations in the absorbing state, retaining only one.

## Details

sim.msm is called repeatedly to produce a simulated trajectory for each individual. The state at each specified observation time is then taken to produce a new column state. The effect of timedependent covariates on the transition intensity matrix for an individual is determined by assuming
that the covariate is a step function which remains constant in between the individual's observation times. If the subject enters an absorbing state, then only the first observation in that state is kept in the data frame. Rows corresponding to future observations are deleted. The entry times into states given in death are assumed to be known exactly.

## Value

A data frame with columns,
subject Subject identification indicators
time Observation times
state $\quad$ Simulated (true) state at the corresponding time
obs Observed outcome at the corresponding time, if ematrix or hmodel was supplied
keep Row numbers of the original data. Useful when drop.absorb=TRUE, to show which rows were not dropped
plus any supplied covariates.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

sim.msm

## Examples

```
### Simulate 100 individuals with common observation times
sim.df <- data.frame(subject = rep(1:100, rep(13,100)), time = rep(seq(0, 24, 2), 100))
qmatrix <- rbind(c(-0.11, 0.1, 0.01),
    c(0.05, -0.15, 0.1 ),
    c(0.02, 0.07, -0.09))
simmulti.msm(sim.df, qmatrix)
```

sojourn.msm Mean sojourn times from a multi-state model

## Description

Estimate the mean sojourn times in the transient states of a multi-state model and their confidence limits.

## Usage

sojourn.msm(x, covariates="mean", ci=c("delta","normal","bootstrap","none"), $\mathrm{cl}=0.95, \mathrm{~B}=1000$ )

## Arguments

B Number of bootstrap replicates, or number of normal simulations from the dis-

X
covariates
ci

A fitted multi-state model, as returned by msm.
The covariate values at which to estimate the mean sojourn times. This can either be:
the string "mean", denoting the means of the covariates in the data (this is the default),
the number 0 , indicating that all the covariates should be set to zero,
a list of values, with optional names. For example,
list $(60,1)$, where the order of the list follows the order of the covariates originally given in the model formula, or a named list, e.g.
list (age $=60$, sex $=1$ )
ci If "delta" (the default) then confidence intervals are calculated by the delta method, or by simple transformation of the Hessian in the very simplest cases.
If "normal", then calculate a confidence interval by simulating $B$ random vectors from the asymptotic multivariate normal distribution implied by the maximum likelihood estimates (and covariance matrix) of the log transition intensities and covariate effects, then transforming.
If "bootstrap" then calculate a confidence interval by non-parametric bootstrap refitting. This is 1-2 orders of magnitude slower than the "normal" method, but is expected to be more accurate. See boot.msm for more details of bootstrapping in $\mathbf{~ m s m}$.
cl Width of the symmetric confidence interval to present. Defaults to 0.95 . tribution of the MLEs

## Details

The mean sojourn time in a transient state $r$ is estimated by $-1 / q_{r r}$, where $q_{r r}$ is the $r$ th entry on the diagonal of the estimated transition intensity matrix.
A continuous-time Markov model is fully specified by the mean sojourn times and the probability that each state is next (pnext.msm). This is a more intuitively meaningful description of a model than the transition intensity matrix (qmatrix.msm).
Time dependent covariates, or time-inhomogeneous models, are not supported. This would require the mean of a piecewise exponential distribution, and the package author is not aware of any general analytic form for that.

## Value

A data frame with components:

| estimates | Estimated mean sojourn times in the transient states. |
| :--- | :--- |
| SE | Corresponding standard errors. |
| L | Lower confidence limits. |
| U | Upper confidence limits. |

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

```
msm, qmatrix.msm, deltamethod
```

```
statetable.msm Table of transitions
```


## Description

Calculates a frequency table counting the number of times each pair of states were observed in successive observation times. This can be a useful way of summarising multi-state data.

## Usage

statetable.msm(state, subject, data=NULL)

## Arguments

state Observed states, assumed to be ordered by time within each subject.
subject Subject identification numbers corresponding to state. If not given, all observations are assumed to be on the same subject.
data An optional data frame in which the variables represented by subject and state can be found.

## Details

If the data are intermittently observed (panel data) this table should not be used to decide what transitions should be allowed in the $Q$ matrix, which works in continuous time. This function counts the transitions between states over a time interval, not in real time. There can be observed transitions between state $r$ and $s$ over an interval even if $q_{r s}=0$, because the process may have passed through one or more intermediate states in the middle of the interval.

## Value

A frequency table with starting states as rows and finishing states as columns.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

crudeinits.msm

## Examples

```
## Heart transplant data
data(cav)
## 148 deaths from state 1, 48 from state 2 and 55 from state 3.
statetable.msm(state, PTNUM, data=cav)
```

surface.msm Explore the likelihood surface

## Description

Plot the log-likelihood surface with respect to two parameters.

## Usage

```
surface.msm(x, params=c(1,2), np=10, type=c("contour","filled.contour","persp","image"),
                point=NULL, xrange=NULL, yrange=NULL,...)
    \#\# S3 method for class 'msm'
    contour ( \(\mathrm{x}, \mathrm{}.\). )
    \#\# S3 method for class 'msm'
    persp(x, ...)
    \#\# S3 method for class 'msm'
    image ( \(x, . .\). )
```


## Arguments

x
params
np
type
"contour" Contour plot, using the R function contour.
"filled. contour" Solid-color contour plot, using the R function filled.contour.
"persp" Perspective plot, using the R function persp.
"image" Grid color plot, using the R function image.
point Vector of length $n$, where $n$ is the number of parameters in the model, including the parameters that will be varied here. This specifies the point at which to fix
the likelihood. By default, this is the maximum likelihood estimates stored in the fitted model $\mathrm{x}, \mathrm{x}$ \$estimates.
xrange $\quad$ Range to plot for the first varied parameter. Defaults to plus and minus two standard errors, obtained from the Hessian at the maximum likelihood estimate.
yrange $\quad$ Range to plot for the second varied parameter. Defaults to plus and minus two standard errors, obtained from the Hessian at the maximum likelihood estimate.
... Further arguments to be passed to the plotting function.

## Details

Draws a contour or perspective plot. Useful for diagnosing irregularities in the likelihood surface. If you want to use these plots before running the maximum likelihood estimation, then just run msm with all estimates fixed at their initial values.
contour.msm just calls surface.msm with type = "contour".
persp.msm just calls surface.msm with type = "persp".
image.msm just calls surface.msm with type = "image".
As these three functions are methods of the generic functions contour, persp and image, they can be invoked as contour $(x)$, persp( $x$ ) or image $(x)$, where $x$ is a fitted msm object.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## See Also

msm, contour, filled. contour, persp, image.
tnorm Truncated Normal distribution

## Description

Density, distribution function, quantile function and random generation for the truncated Normal distribution with mean equal to mean and standard deviation equal to sd before truncation, and truncated on the interval [lower, upper].

## Usage

```
dtnorm(x, mean=0, sd=1, lower=-Inf, upper=Inf, log = FALSE)
ptnorm(q, mean=0, sd=1, lower=-Inf, upper=Inf,
    lower.tail = TRUE, log.p = FALSE)
qtnorm(p, mean=0, sd=1, lower=-Inf, upper=Inf,
    lower.tail = TRUE, log.p = FALSE)
rtnorm(n, mean=0, sd=1, lower=-Inf, upper=Inf)
```


## Arguments

```
    x,q vector of quantiles.
    p vector of probabilities.
    n number of observations. If length (n) > 1, the length is taken to be the number
        required.
    mean vector of means.
    sd vector of standard deviations.
    lower lower truncation point.
    upper upper truncation point.
    log logical; if TRUE, return log density or log hazard.
    log.p logical; if TRUE, probabilities p are given as }\operatorname{log}(p)
    lower.tail logical; if TRUE (default), probabilities are P[X <= x], otherwise, P[X > x].
```


## Details

The truncated normal distribution has density

$$
f(x, \mu, \sigma)=\phi(x, \mu, \sigma) /(\Phi(u, \mu, \sigma)-\Phi(l, \mu, \sigma))
$$

for $l<=x<=u$, and 0 otherwise.
$\mu$ is the mean of the original Normal distribution before truncation,
$\sigma$ is the corresponding standard deviation,
$u$ is the upper truncation point,
$l$ is the lower truncation point,
$\phi(x)$ is the density of the corresponding normal distribution, and
$\Phi(x)$ is the distribution function of the corresponding normal distribution.
If mean or sd are not specified they assume the default values of 0 and 1 , respectively.
If lower or upper are not specified they assume the default values of -Inf and Inf, respectively, corresponding to no lower or no upper truncation.
Therefore, for example, dtnorm( $x$ ), with no other arguments, is simply equivalent to dnorm( $x$ ).
Only rtnorm is used in the msm package, to simulate from hidden Markov models with truncated normal distributions. This uses the rejection sampling algorithms described by Robert (1995).
These functions are merely provided for completion, and are not optimized for numerical stability or speed. To fit a hidden Markov model with a truncated Normal response distribution, use a hmmTNorm constructor. See the hmm-dists help page for further details.

## Value

dtnorm gives the density, ptnorm gives the distribution function, qtnorm gives the quantile function, and $r$ tnorm generates random deviates.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## References

Robert, C. P. Simulation of truncated normal variables. Statistics and Computing (1995) 5, 121-125

## See Also

dnorm

## Examples

```
x <- seq(50, 90, by=1)
plot(x, dnorm(x, 70, 10), type="l", ylim=c(0,0.06)) ## standard Normal distribution
lines(x, dtnorm(x, 70, 10, 60, 80), type="l") ## truncated Normal distribution
```

totlos.msm Total length of stay, or expected number of visits

## Description

Estimate the expected total length of stay, or the expected number of visits, in each state, for an individual in a given period of evolution of a multi-state model.

## Usage

```
totlos.msm(x, start=1, end=NULL, fromt=0, tot=Inf, covariates="mean",
            piecewise.times=NULL, piecewise.covariates=NULL,
            num.integ=FALSE, discount=0, env=FALSE,
            ci=c("none","normal","bootstrap"), cl=0.95, B=1000,
            cores=NULL, ...)
envisits.msm(x, start=1, end=NULL, fromt=0, tot=Inf, covariates="mean",
            piecewise.times=NULL, piecewise.covariates=NULL,
            num.integ=FALSE, discount=0,
            ci=c("none","normal","bootstrap"), cl=0.95, B=1000,
            cores=NULL, ...)
```


## Arguments

$x \quad$ A fitted multi-state model, as returned by msm.
start Either a single number giving the state at the beginning of the period, or a vector of probabilities of being in each state at this time.
end $\quad$ States to estimate the total length of stay (or number of visits) in. Defaults to all states. This is deprecated, since with the analytic solution (see "Details") it doesn't save any computation to only estimate for a subset of states.
fromt Time from which to estimate. Defaults to 0 , the beginning of the process.

| tot | Time up to which the estimate is made. Defaults to infinity, giving the expected time spent in or number of visits to the state until absorption. However, the calculation will be much more efficient if a finite (potentially large) time is specified: see the "Details" section. For models without an absorbing state, $t$ must be specified. |
| :---: | :---: |
| covariates | The covariate values to estimate for. This can either be: |
|  | the string "mean", denoting the means of the covariates in the data (this is the default), |
|  | the number 0 , indicating that all the covariates should be set to zero, |
|  | or a list of values, with optional names. For example |
|  | list (60,1) |
|  | where the order of the list follows the order of the covariates originally given in the model formula, or a named list, |
|  | list (age $=60$, sex $=1$ ) |
| piecewise.times |  |
|  | Times at which piecewise-constant intensities change. See pmatrix. piecewise.msm for how to specify this. This is only required for time-inhomogeneous models specified using explicit time-dependent covariates, and should not be used for models specified using "pci". |
| piecewise.covariates |  |
|  | Covariates on which the piecewise-constant intensities depend. See pmatrix.piecewise.msm for how to specify this. |
| num.integ | Use numerical integration instead of analytic solution (see below). |
| discount | Discount rate in continuous time. |
| env | Supplied to totlos.msm. If TRUE, return the expected number of visits to each state. If FALSE, return the total length of stay in each state. envisits.msm simply calls totlos.msm with env=TRUE. |
| ci | If "normal", then calculate a confidence interval by simulating B random vectors from the asymptotic multivariate normal distribution implied by the maximum likelihood estimates (and covariance matrix) of the log transition intensities and covariate effects, then calculating the total length of stay for each replicate. |
|  | If "bootstrap" then calculate a confidence interval by non-parametric bootstrap refitting. This is 1-2 orders of magnitude slower than the "normal" method, but is expected to be more accurate. See boot.msm for more details of bootstrapping in msm. |
|  | If "none" (the default) then no confidence interval is calculated. |
| cl | Width of the symmetric confidence interval, relative to 1 |
| B | Number of bootstrap replicates |
| cores | Number of cores to use for bootstrapping using parallel processing. See boot.msm for more details. |
|  | Further arguments to be passed to the integrate function to control the numerical integration. |

## Details

The expected total length of stay in state $j$ between times $t_{1}$ and $t_{2}$, from the point of view of an individual in state $i$ at time 0 , is defined by the integral from $t_{1}$ to $t_{2}$ of the $i, j$ entry of the transition probability matrix $P(t)=\operatorname{Exp}(t Q)$, where $Q$ is the transition intensity matrix.
The corresponding expected number of visits to state $j$ (excluding the stay in the current state at time 0 ) is $\sum_{i!=j} T_{i} Q_{i, j}$, where $T_{i}$ is the expected amount of time spent in state $i$.
More generally, suppose that $\pi_{0}$ is the vector of probabilities of being in each state at time 0 , supplied in start, and we want the vector $\mathbf{x}$ giving the expected lengths of stay in each state. The corresponding integral has the following solution (van Loan 1978; van Rosmalen et al. 2013)

$$
\mathbf{x}=\left[\begin{array}{ll}
1 & \mathbf{0}_{K}
\end{array}\right] \operatorname{Exp}\left(t Q^{\prime}\right)\left[\begin{array}{c}
\mathbf{0}_{K} \\
I_{K}
\end{array}\right]
$$

where

$$
Q^{\prime}=\left[\begin{array}{ll}
0 & \pi_{0} \\
\mathbf{0}_{K} & Q-r I_{K}
\end{array}\right]
$$

$\pi_{0}$ is the row vector of initial state probabilities supplied in start, $\mathbf{0}_{K}$ is the row vector of K zeros, $r$ is the discount rate, $I_{K}$ is the $\mathrm{K} \times \mathrm{K}$ identity matrix, and $\operatorname{Exp}$ is the matrix exponential.

Alternatively, the integrals can be calculated numerically, using the integrate function. This may take a long time for models with many states where $P(t)$ is expensive to calculate. This is required where tot = Inf, since the package author is not aware of any analytic expression for the limit of the above formula as $t$ goes to infinity.
With the argument num. integ=TRUE, numerical integration is used even where the analytic solution is available. This facility is just provided for checking results against versions 1.2.4 and earlier, and will be removed eventually. Please let the package maintainer know if any results are different.
For a model where the individual has only one place to go from each state, and each state is visited only once, for example a progressive disease model with no recovery or death, these are equal to the mean sojourn time in each state. However, consider a three-state health-disease-death model with transitions from health to disease, health to death, and disease to death, where everybody starts healthy. In this case the mean sojourn time in the disease state will be greater than the expected length of stay in the disease state. This is because the mean sojourn time in a state is conditional on entering the state, whereas the expected total time diseased is a forecast for a healthy individual, who may die before getting the disease.
In the above formulae, $Q$ is assumed to be constant over time, but the results generalise easily to piecewise-constant intensities. This function automatically handles models fitted using the pci option to msm. For any other inhomogeneous models, the user must specify piecewise.times and piecewise.covariates arguments to totlos.msm.

## Value

A vector of expected total lengths of stay (totlos.msm), or expected number of visits (envisits.msm), for each transient state.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## References

C. van Loan (1978). Computing integrals involving the matrix exponential. IEEE Transactions on Automatic Control 23(3)395-404.
J. van Rosmalen, M. Toy and J.F. O'Mahony (2013). A mathematical approach for evaluating Markov models in continuous time without discrete-event simulation. Medical Decision Making 33:767-779.

## See Also

sojourn.msm, pmatrix.msm, integrate, boot.msm.

$$
\text { transient.msm } \quad \text { Transient and absorbing states }
$$

## Description

Returns the transient and absorbing states of either a fitted model or a transition intensity matrix.

## Usage

transient.msm(x=NULL, qmatrix=NULL)
absorbing.msm(x=NULL, qmatrix=NULL)

## Arguments

x
qmatrix A transition intensity matrix. The diagonal is ignored and taken to be minus the sum of the rest of the row.

## Value

A vector of the ordinal indices of the transient or absorbing states.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

```
    updatepars.msm updatepars.msm
```


## Description

Update the maximum likelihood estimates in a fitted model object. Developer use only.

## Usage

updatepars.msm(x, pars)

## Arguments

$x \quad$ A fitted multi-state model object, as returned by msm.
pars Vector of new parameters, in their untransformed real-line parameterisations, to substitute for the maximum likelihood estimates corresponding to those in the estimates component of the fitted model object (msm. object). The order of the parameters is documented in msm, argument fixedpars.

## Value

An updated msm model object with the updated maximum likelihood estimates, but with the covariances / standard errors unchanged.
Point estimates from output functions such as qmatrix.msm, pmatrix.msm, or any related function, can then be evaluated with the new parameters, and at arbitrary covariate values.
This function is used, for example, when computing confidence intervals from pmatrix.msm, and related functions, using the ci="normal" method.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

$$
\begin{array}{ll}
\text { viterbi.msm } & \begin{array}{l}
\text { Calculate the probabilities of underlying states and the most likely } \\
\text { path through them }
\end{array}
\end{array}
$$

## Description

For a fitted hidden Markov model, or a model with censored state observations, the Viterbi algorithm recursively constructs the path with the highest probability through the underlying states. The probability of each hidden state is also computed for hidden Markov models.

## Usage

viterbi.msm(x, normboot=FALSE)

## Arguments

X
normboot If TRUE, then before running the algorithm, the maximum likelihood estimates of the model parameters are replaced by an alternative set of parameters drawn randomly from the asymptotic multivariate normal distribution of the MLEs.

## Value

A data frame with columns:
subject $=$ subject identification numbers
time $=$ times of observations
observed $=$ corresponding observed states
fitted $=$ corresponding fitted states found by Viterbi recursion. If the model is not a hidden Markov model and there are no censored state observations, this is just the observed states.
For hidden Markov models, an additional matrix pstate is also returned inside the data frame, giving the probability of each hidden state at each point, conditionally on all the data. This is computed by the forward/backward algorithm.

## Author(s)

C. H. Jackson [chris.jackson@mrc-bsu.cam.ac.uk](mailto:chris.jackson@mrc-bsu.cam.ac.uk)

## References

Durbin, R., Eddy, S., Krogh, A. and Mitchison, G. Biological sequence analysis, Cambridge University Press, 1998.

## See Also

msm

## Index

## * Delta method

deltamethod, 13

* Goodness of fit
prevalence.msm, 81
* Matrix exponential

MatrixExp, 33

* Measurement error
medists, 35
* Simulation
sim.msm, 95
simfitted.msm, 96
simmulti.msm, 97
* Survival
plot.msm, 67
* datagen
sim.msm, 95
simmulti.msm, 97
* datasets
aneur, 5
bos, 9
cav, 10
fev, 23
psor, 86
* distribution

2phase, 3
hmm-dists, 25
hmmMV, 28
medists, 35
pexp, 65
qgeneric, 88
tnorm, 103

* math
deltamethod, 13
MatrixExp, 33
* models
boot.msm, 6
coef.msm, 11
crudeinits.msm, 12
draic.msm, 15
efpt.msm, 18
ematrix.msm, 21
hazard.msm, 24
logLik.msm, 32
lrtest.msm, 33
model.frame.msm, 37
msm, 39
msm. form.qoutput, 52
msm. summary, 55
odds.msm, 58
pearson.msm, 61
phasemeans.msm, 66
plot.msm, 67
plot.prevalence.msm, 68
plot.survfit.msm, 70
plotprog.msm, 72
pmatrix.msm, 73
pmatrix. piecewise.msm, 75
pnext.msm, 77
ppass.msm, 79
prevalence.msm, 81
print.msm, 84
printold.msm, 85
qmatrix.msm, 89
qratio.msm, 92
scoreresid.msm, 94
simfitted.msm, 96
sojourn.msm, 99
statetable.msm, 101
surface.msm, 102
totlos.msm, 105
transient.msm, 108
updatepars.msm, 109
viterbi.msm, 109
2phase, 3
absorbing.msm (transient.msm), 108
AIC, 15, 32, 33
aneur, 5
boot.msm, 6, 7, 14, 19-21, 66, 74-76, 78-80, $82,90,92,93,100,106,108$
bos, 9
bos3 (bos), 9
bos 4 (bos), 9
cav, 10
cmodel. object, 11, 54
coef.msm, 11
contour, 102, 103
contour.msm (surface.msm), 102
coxph, 56
crudeinits.msm, 12, 40, 101
d2phase, 48, 67
d2phase (2phase), 3
dbeta, 26
dbinom, 26
deltamethod, 7, 13, 91, 101
deriv, 14
dexp, 26, 66
dgamma, 26
dmenorm (medists), 35
dmeunif (medists), 35
dnbinom, 26
dnorm, 37, 105
dpexp (pexp), 65
dpois, 26
draic.msm, 15
drlcv.msm (draic.msm), 15
dt, 27
dtnorm, 37
dtnorm (tnorm), 103
dunif, 37
dweibull, 26
ecmodel.object, 18, 54
efpt.msm, 18, 80
ematrix.msm, 7, 8, 21, 22, 53, 91
emodel.object, 22, 32, 54, 91
envisits.msm, 106, 107
envisits.msm (totlos.msm), 105
expm, 34
fev, 23
filled. contour, 102, 103
flexsurvreg, 56
format, 52, 84, 85
h2phase (2phase), 3
hazard.msm, 24, 55, 56, 59
hmm-dists, 25
hmmBeta (hmm-dists), 25
hmmBetaBinom (hmm-dists), 25
hmmBinom, 60
hmmBinom (hmm-dists), 25
hmmCat, 41
hmmCat (hmm-dists), 25
hmmExp (hmm-dists), 25
hmmGamma (hmm-dists), 25
hmmIdent, 41, 60
hmmIdent (hmm-dists), 25
hmmLNorm (hmm-dists), 25
hmmMETNorm, 36
hmmMETNorm (hmm-dists), 25
hmmMEUnif, 36
hmmMEUnif (hmm-dists), 25
hmmMV, 27, 28, 28, 39, 41
hmmNBinom (hmm-dists), 25
hmmNorm (hmm-dists), 25
hmmPois (hmm-dists), 25
hmmT (hmm-dists), 25
hmmTNorm, 104
hmmTNorm (hmm-dists), 25
hmmUnif (hmm-dists), 25
hmmWeibull (hmm-dists), 25
hmodel. object, 22, 23, 30, 54, 91
image, 102, 103
image.msm (surface.msm), 102
integrate, 106-108
lines, 68
lines.survfit, 71, 72
load, 7
logLik.msm, 17, 32, 33
lrtest.msm, 32, 33
makeCluster, 7, 16
match.call, 53
MatrixExp, 19, 33, 34, 74, 76, 80
medists, 27,35
model.frame, $38,54,93$
model.frame.msm, 37, 54
model.matrix, 38, 93
model.matrix.msm, 18, 54, 87
model.matrix.msm (model.frame.msm), 37
msm, 6, 7, 11-13, 16, 18, 19, 21-25, 27-34, 38, 39, 51-56, 59-64, 66-68, 70, 73-75,
$77,79,81,84,85,87,89-94,96-98$, 100-103, 105, 107-110
msm. form.eoutput, 85
msm.form.eoutput (msm.form.qoutput), 52
msm.form.qoutput, 51, 52, 85
msm. object, $11,18,23,32,51,53,60,87,91$, 109
msm. summary, 55
msm2Surv, 56, 57
msprep, 58
nlm, 54
odds.msm, 25, 58
optim, 49, 50, 54, 60
p2phase (2phase), 3
par, 68, 69, 71, 72
paramdata, 87
paramdata.object, 55, 59
pearson.msm, 61, 97
persp, 102, 103
persp.msm (surface.msm), 102
pexp, 65
phasemeans.msm, 66
plot, 68, 69, 71, 72
plot.msm, 52,67
plot.prevalence.msm, 68, 72, 83
plot.survfit, 70-73
plot.survfit.msm, 70
plotprog.msm, 72, 73
pmatrix.msm, 6-8, 51, 52, 71, 73, 75, 76, 78, 80, 91, 108, 109
pmatrix.piecewise.msm, $7,8,74,75,75,82$, 106
pmenorm (medists), 35
pmeunif (medists), 35
pnext.msm, 77, 100
pnorm, 88
ppass.msm, 20, 79
ppexp (pexp), 65
prevalence.msm, 7, 8, 55, 56, 63, 64, 69, 81
print.msm, 51-53, 84, 85
print. summary.msm (msm. summary), 55
printnew.msm (print.msm), 84
printold.msm, 85,85
psor, 86
ptnorm (tnorm), 103
q2phase (2phase), 3
qcmodel.object, 54, 87
qgeneric, 88
qmatrix.msm, $7,8,19,22,47,51-53,66,75$, $78,79,81,84,89,90,93,100,101$, 109
qmenorm (medists), 35
qmeunif (medists), 35
qmodel.object, 23, 32, 54, 91
qpexp, 88
qpexp (pexp), 65
qratio.msm, 7, 8, 78, 92
qtnorm (tnorm), 103
r2phase (2phase), 3
recreate.olddata, 93
registerDoParallel, 7, 16
rmenorm (medists), 35
rmeunif (medists), 35
rpexp (pexp), 65
rtnorm (tnorm), 103
save, 7
scoreresid.msm, 64, 94
sim.msm, 66, 95, 97-99
simfitted.msm, 96
simmulti.msm, 52, 96, 97, 97
sojourn.msm, 8, 20, 51, 52, 55, 77, 91, 99, 108
statetable.msm, 13, 101
summary.msm, 24, 25, 52, 84
summary.msm (msm. summary), 55
surface.msm, 102
survfit, 72,73
tnorm, 27, 103
totlos.msm, 7, 8, 20, 80, 105, 106, 107
transient.msm, 108
uniroot, 88
updatepars.msm, 109
viterbi.msm, 83, 109

